



VetAgro Sup



STARTING WITH



Karine Chalvet-Monfray



Mahidol
University
Wisdom of the Land

Faculty of Veterinary Science – Mahidol University – 22-24 July 2015

Pedagogical aims

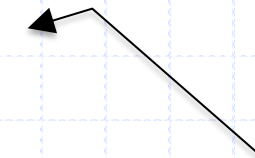
- ◆ Know how import data in 
- ◆ Know how do basic representation
- ◆ Know how do simple statistic tests
- ◆ Know how do calculus of epidemiology
- ◆ Know how learn more on 

Color code concerning code

function()

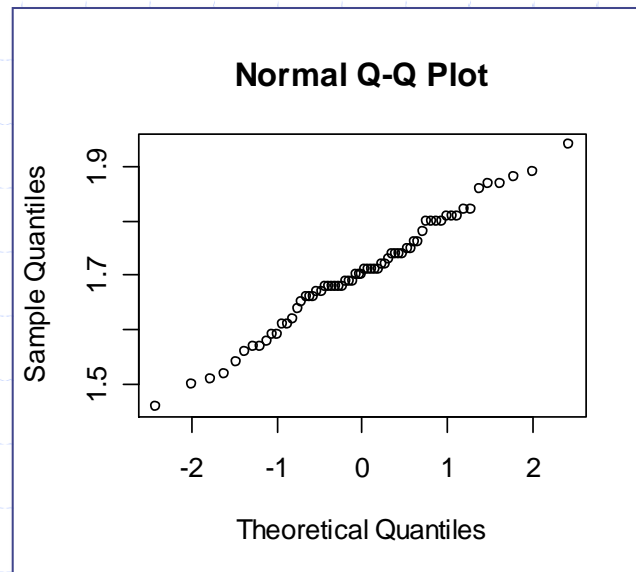
2+2
[1] 4

d1 is a data frame
d1 has 2 column=2 variables



Don't forget
to close
before asking
a new code

Big sample
(N>50)



Color code concerning variables

V1
9.0
0.8
5.9
1.6
7.6
9.9
0.1
3.1

Quantitative variable

V1

Qualitative variable

V1
+
+
-
+
±
-
-
+

2 qualitative variables

V1	V2
	+
	+
	-
	+
	±
	-
	-
	+

V1	V2
	9.0
	0.8
	5.9
	1.6
	7.6
	9.9
	0.1
	3.1

1 qualitative variable
+ 1 quantitative variable

V1	V2
7.4	9.0
1.8	0.8
5.4	5.9
2.0	1.6
0.6	7.6
5.3	9.9
0.7	0.1
4.8	3.1

2 quantitative variables

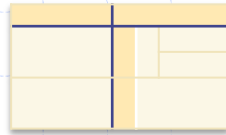
Date
22/06/1965
20/03/1968
30/08/1965
27/09/1965
22/07/1966
05/05/1964
16/01/1964
11/02/1968

Qualitative variable
-> text

Color code concerning plan



Chap. 1



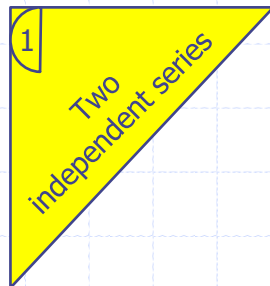
Chap. 4



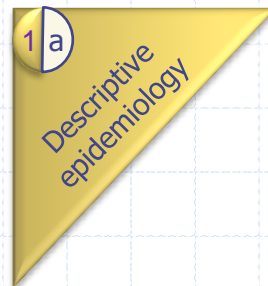
Part 5.1



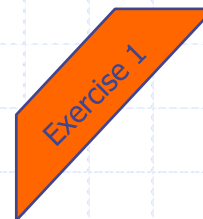
Part 5.2



Basic
card



Perfecting
card



Basic
exercise



Perfecting
exercise

What is ?

◆ Origin

-> R = Language and Software created by Ross Ihaka and Robert Gentleman from the language S (S+: Software) for statistics

◆ Interests

-> worldly used with a lot of help
-> free and open source
-> build by good statisticians and numericians
-> flexibility because using interpreted language
-> increasing tools and field (mathematics, GIS,...)

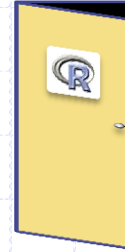
◆ disadvantages

-> less convivial because using language, need help a the beginning

Plan

1. Principle of R language
2. Importing data
3. Simple descriptive statistics -> *Graphics*
4. Simple analytical statistics -> *tests of comparison mean and frequencies*
5. Specific tools for epidemiology
6. Specific tools for clinical study

Plan



1. Principle of R language
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Principle of language

1. R can calculate as a calculator

```
2+2  
[1] 4
```

2. R use functions with arguments [to be specified or optional (set by default)]

function()

```
read.table("cohort.txt", header=TRUE)
```

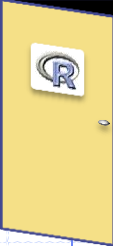
optional argument

argument to be specified

3. The result can be assigned

```
d<-read.table("cohort.txt", header=TRUE)
```

```
read.table("cohort.txt", header=TRUE) -> d
```



Practical on principle of R language

- ◆ Calculate:

$$\sqrt{4}$$

```
4^(0.5)  
[1] 2
```

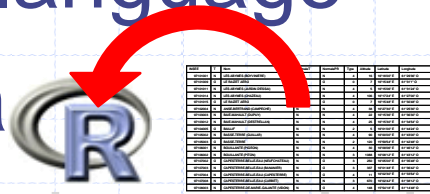
$$\ln(\pi \times (4 \times 10^2 + 543))$$

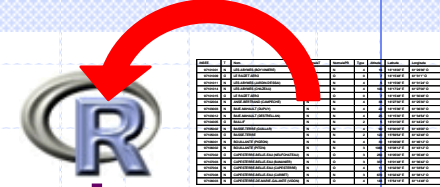
```
log(pi*(4e2+543))  
[1] 7.993796
```

$$e^{2.36-1.96*0.49}$$

```
exp(2.36-1.96*0.49)  
[1] 4.053578
```

Plan

1. Principle of R language
2. Importing data A red curved arrow points from the R logo to a small table of data. The table has multiple columns and rows, representing a dataset.
3. Simple descriptive statistics -> *Graphics*
4. Simple analytical statistics -> *tests of comparison mean and frequencies*
5. Specific tools for epidemiology
6. Specific tools for clinical study



Importing data: one method

1. In Excel save as txt format (tabulation as a separator) the data frame

If there is empty cell write NA for Not Available

2. Read the table in R and assign the result

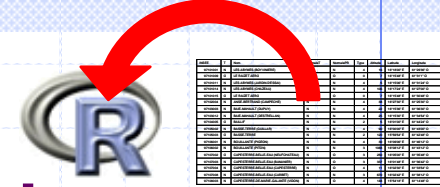
name of
results =
name of
data frame

function which reads a table

```
d<-read.table("cohort.txt",header=T)
```

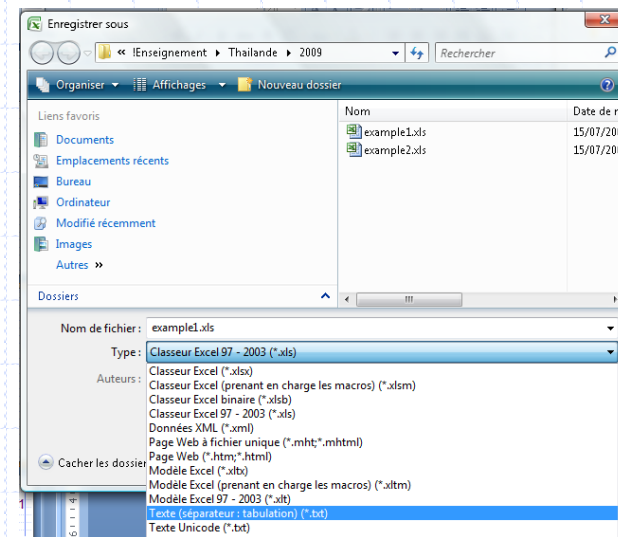
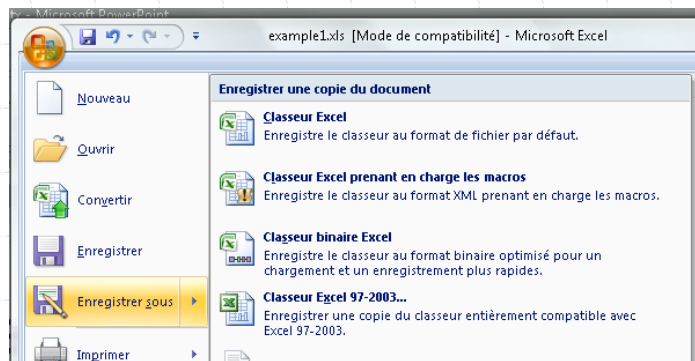
"name of txt file"

the table has header

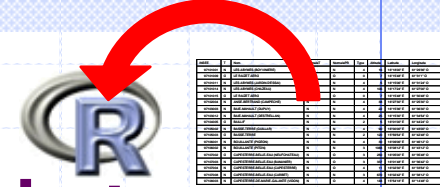


Practical on importing data

Using the file example1.xls, import the data in R with assigning the result to d1.



```
d1<-read.table("exemple1.txt",header=T)
```



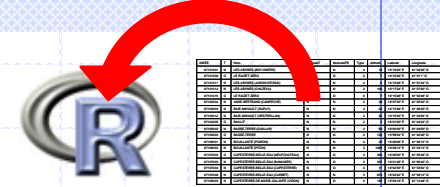
Practical on importing data

Do it the same with example2.xls.

It doesn't work. See data in Excel.

	A	B	C
1	origin	size body	dosage
2	A	1.5	295.7
3	A	1.64	50.5
4	A	1.83	136.6
5	A	1.57	107.1
6	A	1.73	329.9
7	A		418.7
8	A	1.66	209.9
9	A	1.73	140.3
10	A	1.75	36.5

	A	B	C
1	origin	size_body	dosage
2	A	1.5	295.7
3	A	1.64	50.5
4	A	1.83	136.6
5	A	1.57	107.1
6	A	1.73	329.9
7	A	NA	418.7
8	A	1.66	209.9
9	A	1.73	140.3
10	A	1.75	36.5



Importing data: quick visualization 1

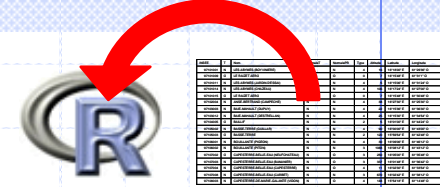
Some R functions are very useful to visualize the data frame: `view()`

`view(d1)`

function which edits
an object

the data frame

	origin	size
1	A	1.64
2	A	1.61
3	A	1.68
4	A	1.87
5	A	1.52
6	A	1.76
7	A	1.69
8	A	1.81
9	A	1.71
10	A	1.80
11	A	1.70



Importing data: quick visualization 2

Some R functions are very useful to visualize the data frame: `str()`

function which displays compactly the structure of an object

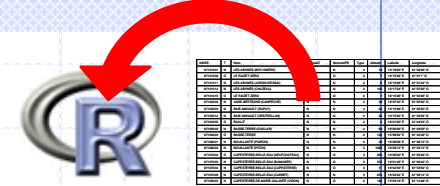
d1 is a data frame
d1 has 65 rows=65 observations
d1 has 2 column=2 variables

`str(d1)`

```
'data.frame':   65 obs. of  2 variables:
 $ origin: Factor w/ 5 levels "A","B","C","D",...: 1 1 1 1 1 1 1 1 1 1 ...
 $ size : num  1.64 1.61 1.68 1.87 1.52 1.76 1.69 1.81 1.71 1.8 ...
```

The first column is named origin. It is a factor variable (qualitative) with 5 levels "A","B",... The first values are A,A,A... (1 is for the first level =A)

The second column is named size. It is a numerical variable (quantitative data). The first values are 1.64,1.61,1.68 ...



Importing data: quick visualization 3

Some R functions are very useful to visualize the data frame: `summary()`

function which produces result summaries

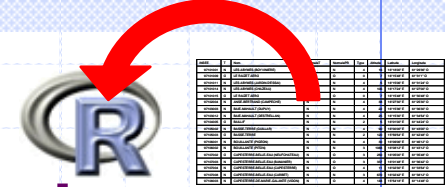
```
summary(d1)
```

origin	size
A:15	Min. :1.460
B:14	1st Qu.:1.660
C:13	Median:1.700
D:12	Mean :1.704
E:11	3rd Qu.:1.760
	Max. :1.940

...

The first variable named origin is qualitative variable. There is 15 observations with level A, 14 obs...

The second variable named size is quantitative variable. The minimum value is 1.46, The first quartile is 1.66...



Practical on importing data

Make a quick visualization with the data of the example 2.

```
view(d2)
```

	origin	size_body	dosage	v
1	A	1.5	295.7	
2	A	1.64	50.5	
3	A	1.83	136.6	
4	A	1.57	107.1	
5	A	1.73	329.9	
6	A	NA	418.7	

```
summary(d2)
```

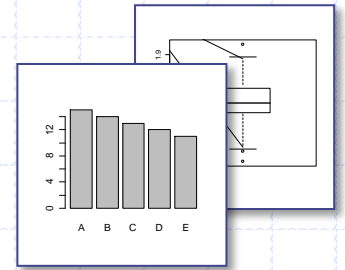
```
origin      size_body      dosage
A:15      Min.   :1.470      Min.    :   6.5
B:14      1st Qu.:1.647      1st Qu.:  60.2
C:13      Median:1.715      Median: 139.6
D:13      Mean   :1.711      Mean    : 224.4
E:10      3rd Qu.:1.765      3rd Qu.: 326.7
          Max.   :1.980      Max.    :1023.6
          NA's   :1.000
```

```
str(d2)
```

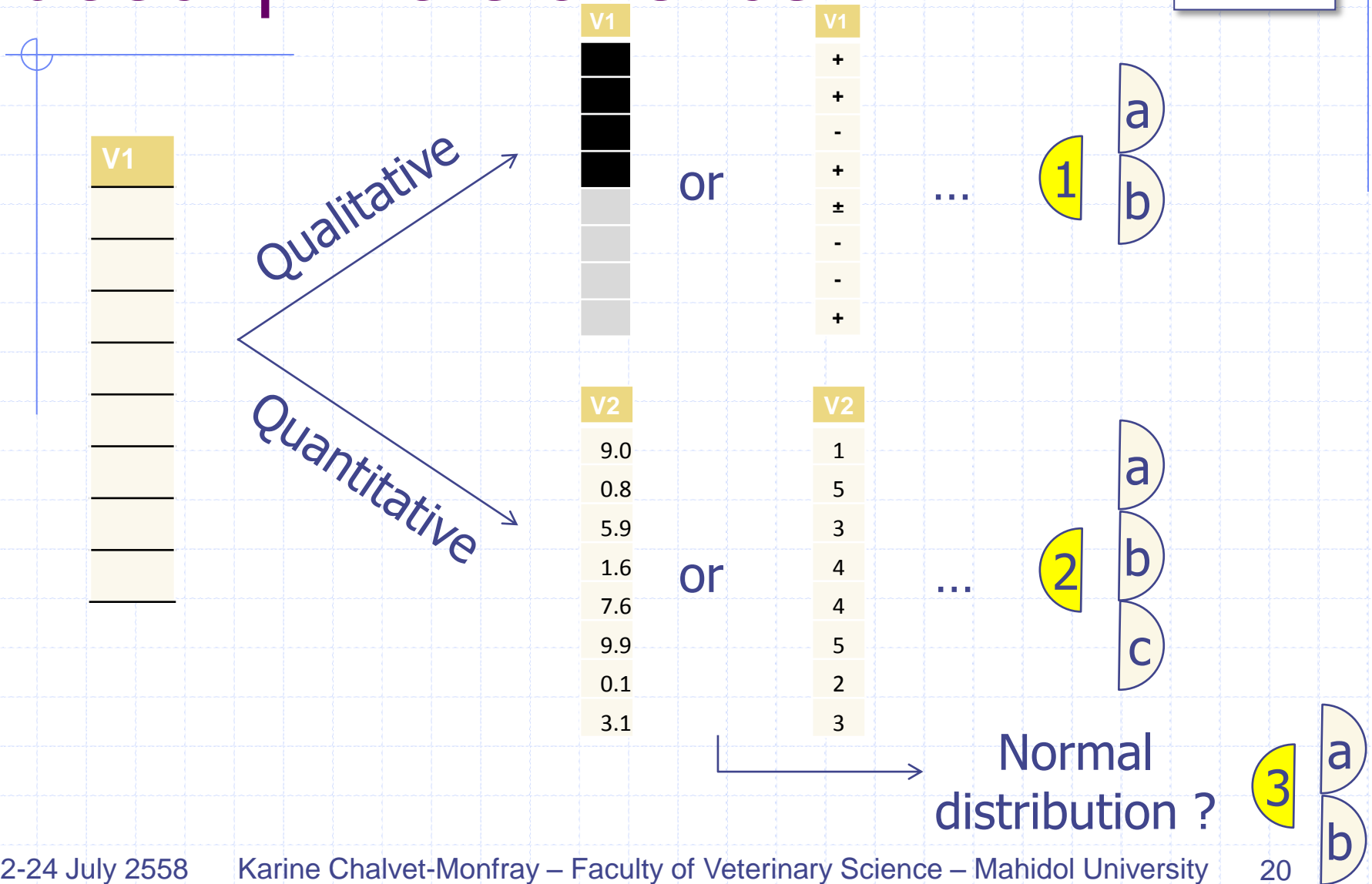
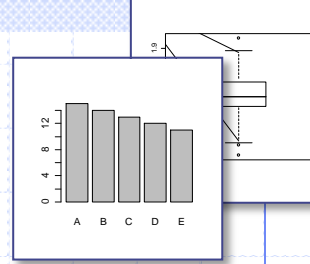
```
'data.frame':   65 obs. of  3 variables:
 $ origin   : Factor w/ 5 levels "A","B","C","D",...: 1 1 1 1 1 1 1 1 1 1 ...
 $ size_body: num  1.5 1.64 1.83 1.57 1.73 NA 1.66 1.73 1.75 1.71 ...
 $ dosage   : num  295.7  50.5 136.6 107.1 329.9 ...
```

Plan

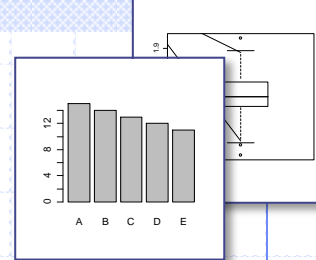
1. Principle of R language
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Introduction for simple descriptive statistics



Simple descriptive statistics for qualitative variable:



`table()` and `barplot()`

function which builds a contingency table

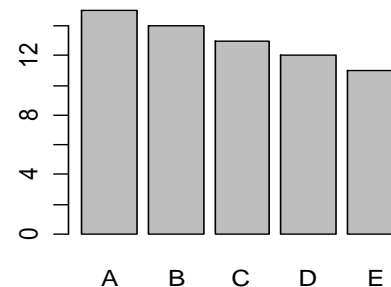
a qualitative variable of the data frame

function which creates a bar plot

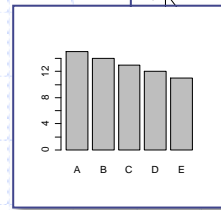
```
table(d1$origin)
```

A	B	C	D	E
15	14	13	12	11

```
barplot(table(d1$origin))
```



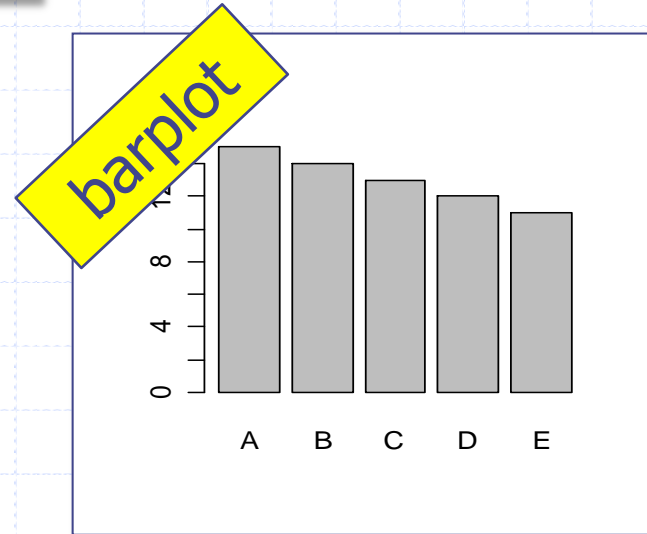
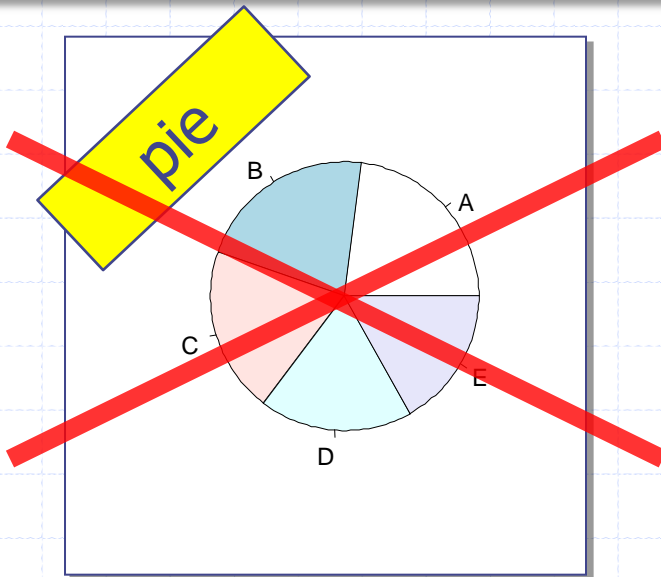
Simple descriptive statistics for qualitative variable:



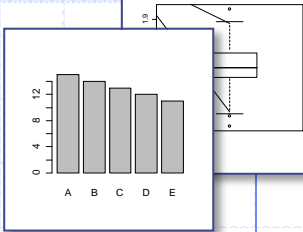
~~pie()~~

function which draws
a pie chart

```
pie(table(d1$origin))
```



Practical on simple descriptive statistics

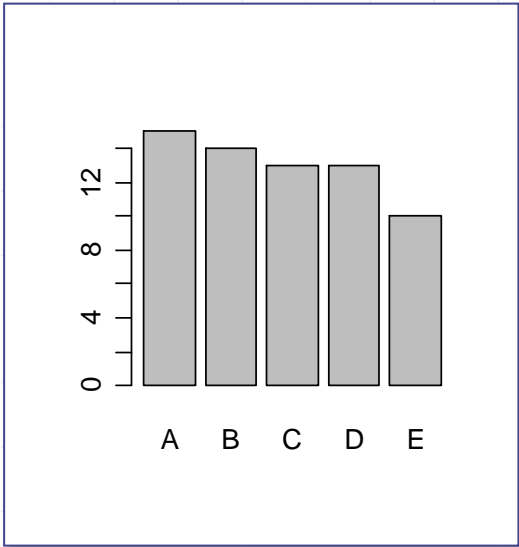


Make a contingency table and a bar plot for the variable “origin” of the example 2.

```
table(d2$origin)
```

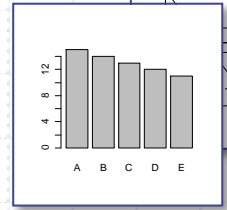
A	B	C	D	E
15	14	13	13	10

```
barplot(table(d2$origin))
```



2a

Simple descriptive statistics for quantitative variable:



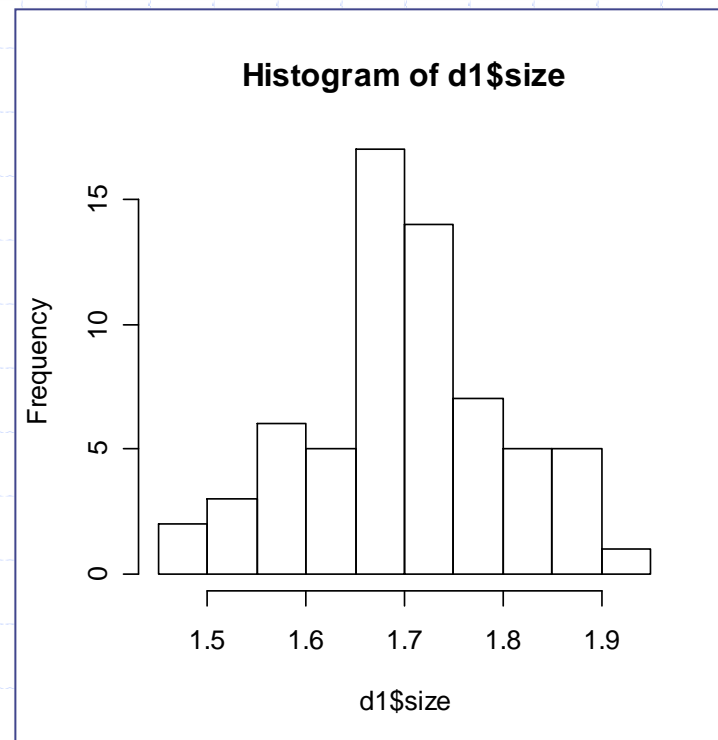
```
hist()
```

function which
computes a histogram

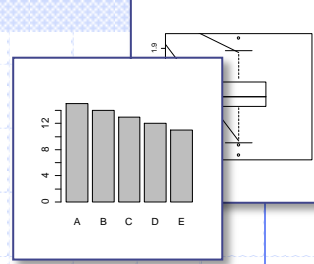
a quantitative variable of
the data frame

```
hist(d1$size)
```

Big sample
($N > 50$)



Simple descriptive statistics for quantitative variable:

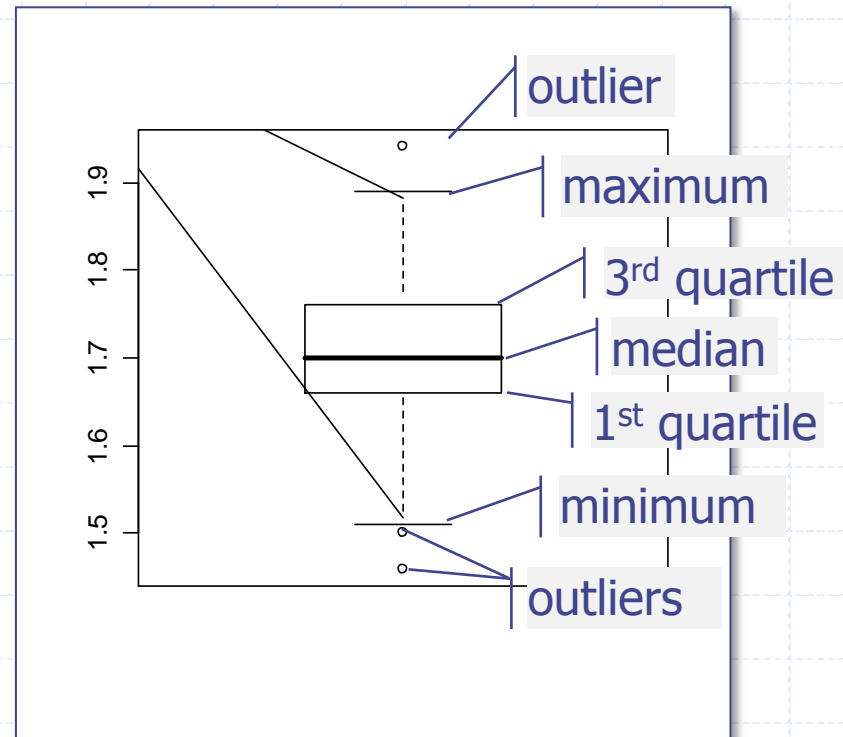


`boxplot()`

function which produces a box-and-whisker plot

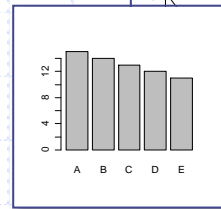
`boxplot(d1$size)`

Not too small sample ($N > 15$)



Maximum and minimum are calculated as the most extreme data points which are no more than 1.5 times the interquartile (3^{rd} quartile - 1^{st} quartile) from the box.

Simple descriptive statistics for quantitative variable:

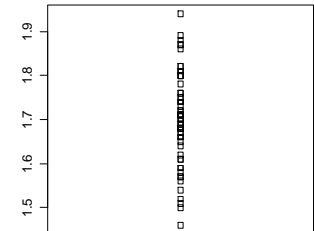


`stripchart()`

function which produces a dot plot (1D scatter plot)

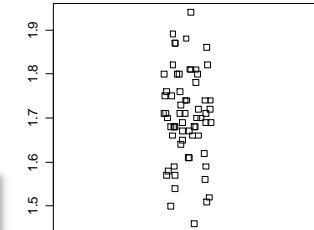
argument for which plot is drawn vertically

```
stripchart(d1$size, vertical=TRUE)
```



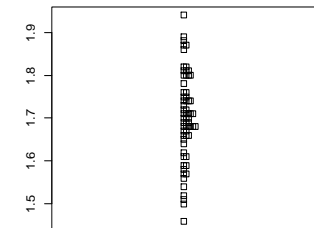
argument for which dots are not overplotted but jittered

```
stripchart(d1$size, vertical=TRUE, method="jitter")
```



```
stripchart(d1$size, vertical=TRUE, method="stack")
```

argument for which dots are not overplotted but stacked



Even for small sample

V2

9.0

0.8

5.9

1.6

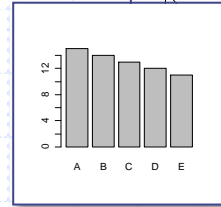
7.6

9.9

0.1

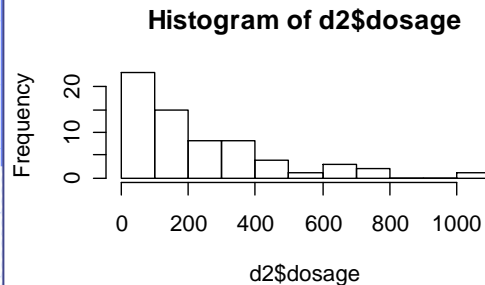
3.1

Practical on simple descriptive statistics

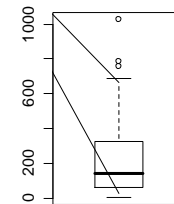


Make a histogram, a box plot and dot plot of the variable “dosage” of the example 2

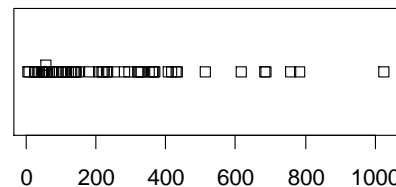
hist(d2\$dosage)



boxplot(d2\$dosage)



stripchart(d2\$dosage, method="stack")



V2

9.0

0.8

5.9

1.6

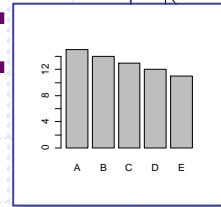
7.6

9.9

0.1

3.1

Simple descriptive statistics: looking for normality:

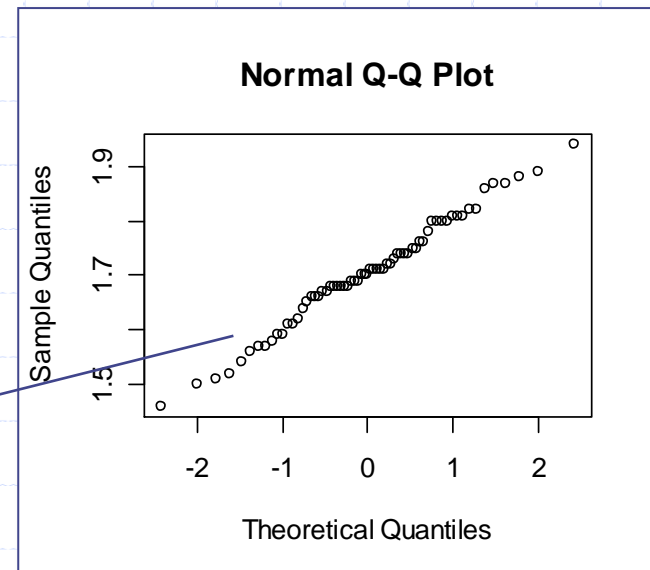


`qqnorm()`

function which produces a
normal QQ plot

`qqnorm(d1$size)`

As the plot looks like a
straight line, the
distribution seems normal



V2

9.0

0.8

5.9

1.6

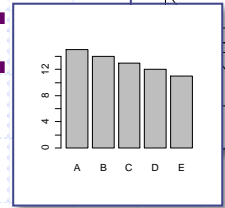
7.6

9.9

0.1

3.1

Simple descriptive statistics: looking for normality:



```
shapiro.test()
```

function which performs the
Shapiro-Wilk test of normality

```
shapiro.test(d1$size)
```

Shapiro-Wilk normality test

data: d1\$size

W = 0.9873, p-value = 0.7448

$p > 0.05$, the normality
hypothesis test is not
rejected;
the distribution seems
normal.

V2

9.0

0.8

5.9

1.6

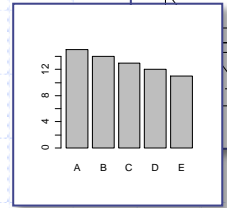
7.6

9.9

0.1

3.1

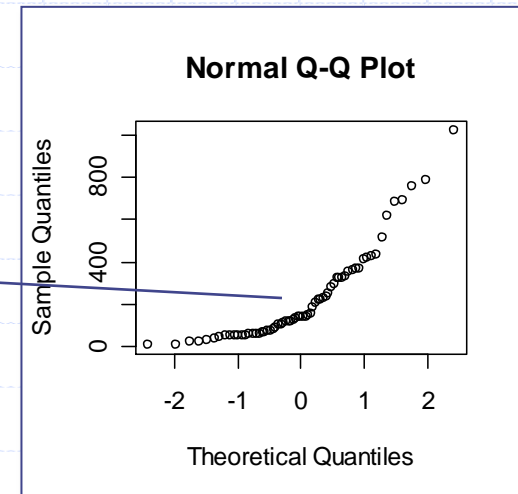
Simple descriptive statistics



Looking for normality of the variable dosage of the example 2.

```
qqnorm(d2$dosage)
```

As the plot doesn't look like a straight line, the distribution is not normal.



```
shapiro.test(d2$dosage)
```

Shapiro-Wilk normality test

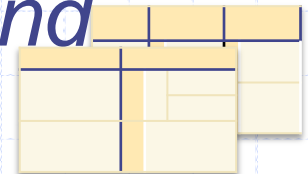
```
data: d2$dosage
```

```
W = 0.8218, p-value = 2.135e-07
```

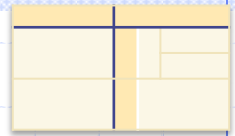
$p < 0.001$, the normality hypothesis is rejected; the distribution is not normal.

Plan

1. Principle of R language
2. Importing data
3. Simple descriptive statistics -> *Graphics*
4. Simple analytical statistics -> *tests of comparison 2 observed mean and frequencies*
5. Specific tools for epidemiology
6. Specific tools for clinical study



Introduction for simple analytical statistics



In order to simplify, we envisaged only to test two series of data.

The series can be:

either independent or paired.

The variable can be:

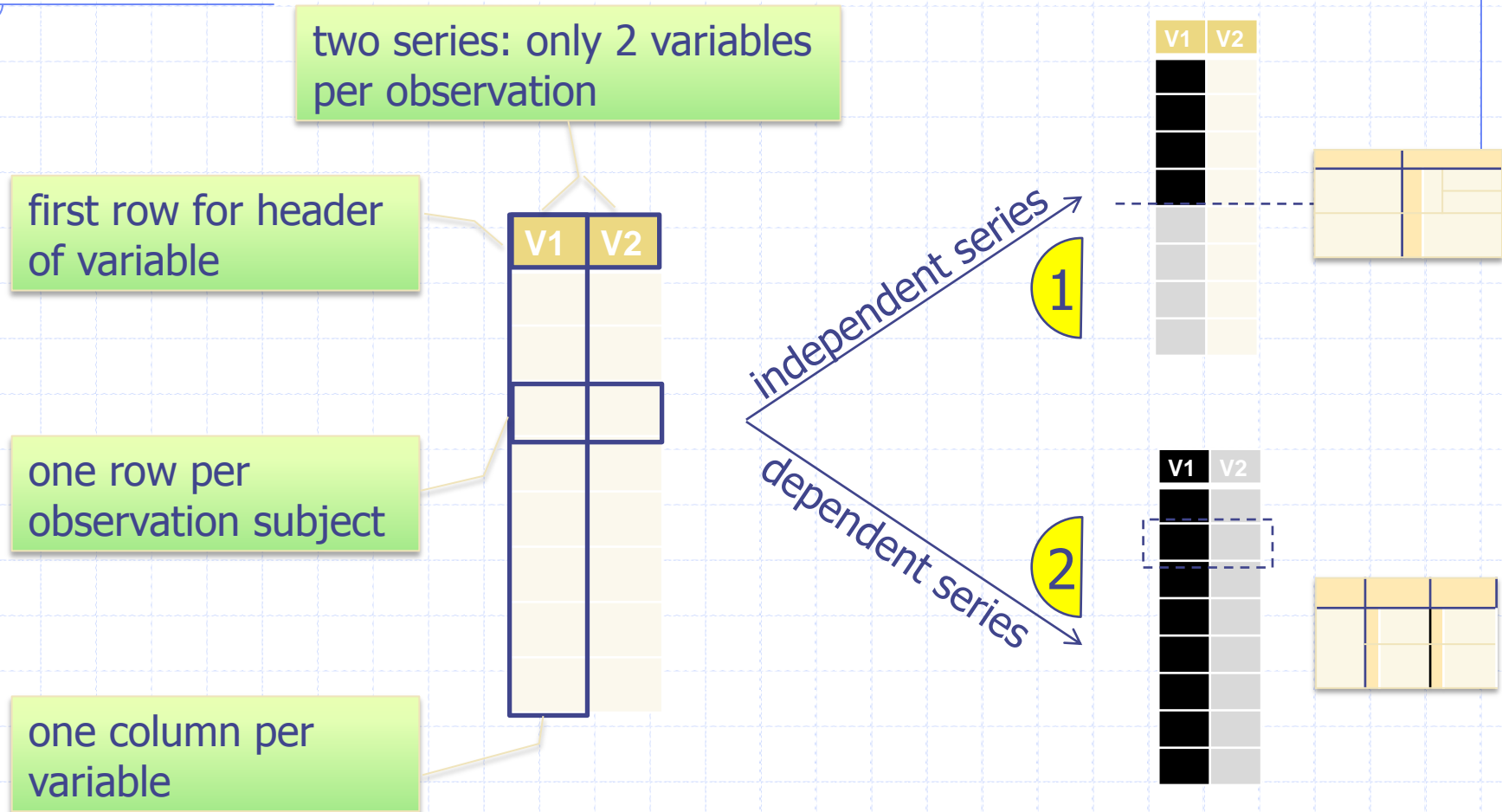
either qualitative or quantitative.

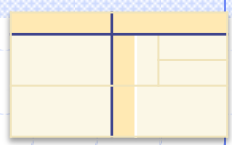
The test for quantitative(s) variable(s) can be:

either parametric or non parametric.

We envisaged R code mainly for raw data.

Introduction for data organization for two series





Different cases for two independent series

Series variable, is dichotomic because there is two series

V1	V2
Black	Yellow
Black	Yellow
Black	Yellow
Black	Yellow
Grey	Yellow
Grey	Yellow
Grey	Yellow
Grey	Yellow

The other variable can be:

- qualitative (dichotomic or not)
- quantitative (normal distribution or not)

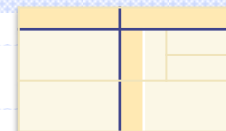
V1	V2
Black	+
Black	+
Black	-
Black	+
Grey	±
Grey	-
Grey	-
Grey	+

V1	V2
Black	9.0
Black	0.8
Black	5.9
Black	1.6
Grey	7.6
Grey	9.9
Grey	0.1
Grey	3.1

a
b

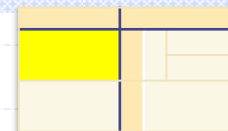
c
d
e
f

Different tests



One dichotomic variable(series) + one qualitative (dichotomic or not) variable	One dichotomic variable (series) + one quantitative (normal or not) variable		
Pearson's Chi-squared Test	parametric tests	F test for two variances	Welch two sample t test
Fisher's exact test only for 2 dichotomic variables			Student two samples t test
	non parametric tests		Mann-Whitney-Wilcoxon rank sum test

The choice between tests for quantitative variable is supposed known



Pearson's Chi-squared test:

```
chisq.test()
```

two qualitative variables

```
chisq.test(d3$exposition, d3$disease)
```

Pearson's Chi-squared test with Yates' continuity correction

data: d3\$exposition and d3\$disease

X-squared = 5.1212, df = 1, p-value = 0.02363

$p < 0.05$, the independence hypothesis test is rejected; the difference is significant.

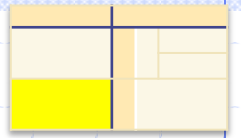


Warning message:

Sample too small:
expected value < 5

V1	V2
	+
	+
	-
	+
	±
	-
	-
	+

1b
Two independent series



Fisher's exact test:

```
fisher.test()
```

2 qualitative variables

```
fisher.test(d3$exposition,d3$disease)
```

Fisher's Exact Test for Count Data

data: d3\$exposition and d3\$disease

p-value = 0.01773

alternative hypothesis: true odds ratio is not equal to 1

95 percent confidence interval:

1.103106 4.395369

sample estimates:

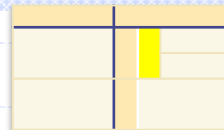
odds ratio

2.178838

Particularly adapted for
small samples (only for 2
dichotomic variables)

$p < 0.05$, the independence hypothesis test is rejected;
the difference is significant.

V1	V2
■	+
■	+
■	-
■	+
■	-
■	-
■	-
■	+



F test for 2 variances:

`var.test()`

a quantitative variable



a qualitative dichotomous variable

```
var.test(d3$response~d3$disease)
```

F test to compare two variances

data: d3\$response by d3\$disease

F = 1.1702, num df = 143, denom df = 54, p-value = 0.5147

alternative hypothesis: true ratio of variances is not equal to 1

95 percent confidence interval:

0.7320177 1.7876995

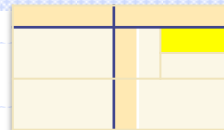
sample estimates:

ratio of variances

1.170242

$p > 0.05$, the hypothesis test H_0 is not rejected;
the difference is not significant.

V1	V2
	9.0
	0.8
	5.9
	1.6
	7.6
	9.9
	0.1
	3.1



Welch two samples t test:

`t.test()`

a quantitative variable



a qualitative dichotomous variable

```
t.test(d3$response~d3$disease)
```

Welch Two Sample t-test

data: d3\$response by d3\$disease

t = -3.4249, df = 105.129, p-value = 0.0008781

alternative hypothesis: true difference in means is not equal to 0

95 percent confidence interval:

-0.7937418 -0.2116779

sample estimates:

mean in group no mean in group yes

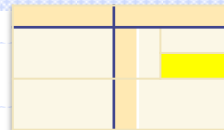
5.110023

<

5.612732

p<0.001, the hypothesis test H0 is rejected; the difference is highly significant.

V1	V2
	9.0
	0.8
	5.9
	1.6
	7.6
	9.9
	0.1
	3.1



Student two samples t test:

t.test()

a quantitative variable



a qualitative dichotomic variable

```
t.test(d3$response~d3$disease, var.equal=TRUE)
```

var.equal=TRUE

Two Sample t-test

data: d3\$response by d3\$disease

t = -3.3062, df = 197, p-value = 0.001124

alternative hypothesis: true difference in means is not equal to 0

95 percent confidence interval:

-0.8025613 -0.2028584

sample estimates:

mean in group no mean in group yes

5.110023

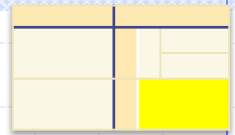


5.612732

p<0.01, the hypothesis test H0 is rejected; the difference is highly significant.

V1	V2
	9.0
	0.8
	5.9
	1.6
	7.6
	9.9
	0.1
	3.1

Mann-Whitney-Wilcoxon rank sum test:



`wilcox.test()`

a quantitative variable



a qualitative dichotomic variable

```
wilcox.test(d3$response~d3$disease)
```

```
Wilcoxon rank sum test with continuity correction
```

```
data: d3$response by d3$disease
```

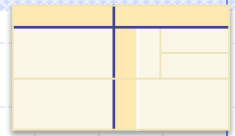
```
W = 2774, p-value = 0.001102
```

```
alternative hypothesis: true location shift is not equal to 0
```

$P < 0.01$, the hypothesis test H_0 is rejected; the difference is highly significant.

Particularly adapted for small samples

V1	V2
	9.0
	0.8
	5.9
	1.6
	7.6
	9.9
	0.1
	3.1



Practical on simple analytical statistics

Test if the two variables “exposition” and “disease” of the example 3 are independent with a Chi² test.

```
chisq.test(d6$exposition,d6$disease)
```

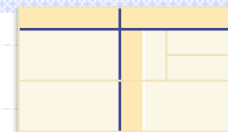
```
Pearson's Chi-squared test with Yates' continuity correction
```

```
data: d6$exposition and d6$disease
```

```
X-squared = 4.5608, df = 1, p-value = 0.03271 S+
```

$p < 0.05$, the independence hypothesis test is rejected; the difference is significant.

Practical on simple analytical statistics



With the same data, test if the variable “dosage” is different according the “exposition” with a student two sample t test. We need to test the homogeneity of the both variances before.

```
var.test(d6$dosage~d6$exposition)
```

F test to compare two variances

data: d6\$dosage by d6\$exposition

F = 0.9378, num df = 100, denom df = 98, p-value = 0.7493 NS

...

```
t.test(d6$dosage~d6$exposition,var.equal=T)
```

Two Sample t-test

data: d6\$dosage by d6\$exposition

t = 2.2108, df = 198, p-value = 0.02820 S+

alternative hypothesis: true difference in means is not equal to 0

95 percent confidence interval:

0.03456140 0.60549861

sample estimates:

mean in group no mean in group yes

5.458515

>

5.138485

Different cases for two dependent series

We will consider only the case where both variables have the same type.

V1	V2

The both variables can be:

- qualitative (dichotomic only)
- quantitative (normal distribution or not)

V1	V2
+	+
+	+
+	-
-	-
-	+
-	+
+	+
-	-

a

V1	V2
6.0	4.4
1.7	8.6
7.1	9.0
3.0	5.0
4.9	4.3
5.2	4.3
0.2	10.0
9.7	6.1

V1-V2
1.6
-6.9
-1.9
-2.0
0.6
0.9
-9.8
3.6

b

c

V1	V2
6.0	4.4
1.7	8.6
7.1	9.0
3.0	5.0
4.9	4.3
5.2	4.3
0.2	10.0
9.7	6.1

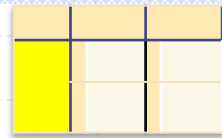
d

e

Different tests

dichotomic variable	One quantitative variable (normal or not)	Two quantitative variables (normal or not)
<div> <div>V1</div> <div>V2</div> <div>+</div><div>+</div> <div>+</div><div>+</div> <div>+</div><div>-</div> <div>-</div><div>-</div> <div>-</div><div>+</div> <div>-</div><div>+</div> <div>+</div><div>+</div> <div>-</div><div>-</div> </div>	<div> <div>V1-V2</div> <div>1.6</div> <div>-6.9</div> <div>-1.9</div> <div>-2.0</div> <div>0.6</div> <div>0.9</div> <div>-9.8</div> <div>3.6</div> </div>	<div> <div>V1</div> <div>V2</div> <div>6.0</div><div>4.4</div> <div>1.7</div><div>8.6</div> <div>7.1</div><div>9.0</div> <div>3.0</div><div>5.0</div> <div>4.9</div><div>4.3</div> <div>5.2</div><div>4.3</div> <div>0.2</div><div>10.0</div> <div>9.7</div><div>6.1</div> </div>
<div> <div>McNemar's Chi-squared Test</div> <div>Particularly adapted for small samples</div> <div>a</div> </div>	<div> <div>parametric tests</div> <div>Paired t test</div> <div>b</div> </div>	<div> <div>parametric tests</div> <div>Pearson's correlation coefficient test</div> <div>d</div> </div>
<div> <div>non parametric tests</div> <div>Wilcoxon signed rank test</div> <div>c</div> </div>	<div> <div>non parametric tests</div> <div>Spearman's rank correlation coefficient test</div> <div>e</div> </div>	

The linear regression is not envisaged here.



McNemar's Chi-squared test:

```
mcnemar.test()
```

two dichotomic variables

```
mcnemar.test(d4$test1,d4$test2)
```

McNemar's Chi-squared test with continuity correction

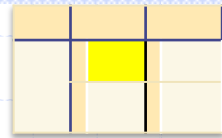
data: d4\$test1 and d4\$test2

McNemar's chi-squared = 32.0727, df = 1, p-value = 1.485e-08

$p < 0.001$, the hypothesis test is rejected;
the difference is highly significant.

Particularly adapted
for small samples

V1	V2
+	+
+	+
+	-
-	-
-	+
-	+
+	+
-	-



Paired t test:

`t.test()`

two quantitative variables

paired=TRUE

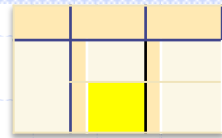
```
t.test(d4$femurRsize,d4$femurLsize,paired=TRUE)
```

Paired t-test

```
data: d5$femurRsize and d5$femurLsize
t = 0.5755, df = 199, p-value = 0.5656
alternative hypothesis: true difference in means is not equal to 0
95 percent confidence interval:
 -0.04974689  0.09074689
sample estimates:
mean of the differences
                0.0205
```

$p > 0.05$, the hypothesis test is not rejected;
the difference is not significant.

V1-V2
1.6
-6.9
-1.9
-2.0
0.6
0.9
-9.8
3.6



Wilcoxon signed rank test:

```
wilcox.test()
```

two quantitative variables

paired=TRUE

```
wilcox.test(d4$femurRsize, d4$femurLsize, paired=TRUE)
```

Wilcoxon signed rank test with continuity correction

data: d5\$femurRsize and d5\$femurLsize

V = 8238, p-value = 0.5966

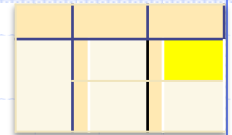
alternative hypothesis: true location shift is not equal to 0

$p > 0.05$, the hypothesis test is not rejected;
the difference is not significant.

Particularly adapted
for small samples

V1-V2
1.6
-6.9
-1.9
-2.0
0.6
0.9
-9.8
3.6

Pearson's correlation coefficient test:



`cor.test()`

two quantitative variables

```
cor.test(d4$size, d4$weightbefore)
```

Pearson's product-moment correlation

data: d5\$size and d5\$weightbefore

t = 9.5205, df = 198, p-value < 2.2e-16

alternative hypothesis: true correlation is not equal to 0

95 percent confidence interval:

0.4571813 0.6486851

sample estimates:

cor

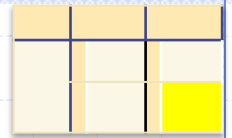
0.5603776

Correlation coefficient r

95% confidence interval of correlation coefficient r

p < 0.001, the hypothesis test is rejected;
the correlation is highly significant.

Spearman's rank correlation coefficient test:



`cor.test()`

two quantitative variables

method="spearman"

```
cor.test(d4$size,d4$weightbefore,method="spearman")
```

Spearman's rank correlation rho

data: d5\$size and d5\$weightbefore
S = 644607, p-value = 4.907e-15

alternative hypothesis: true rho is not equal to 0

sample estimates:

rho
0.5165327 | Correlation coefficient Rho

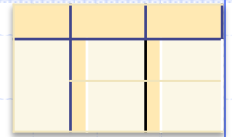
p<0.001, the hypothesis test is rejected;
the correlation is highly significant.

Particularly adapted
for small samples

Warning message:

In `cor.test.default(d5$size, d5$weightbefore, method = "spearman")`:

Impossible de calculer les p-values exactes avec des ex-aequos



Practical on simple analytical statistics

Using data of example 4, test if the two variables “testA” and “testB” realized on the same subject are different with a McNemar Chi² test.

```
mcnemar.test(d7$testA,d7$testB)
```

McNemar's Chi-squared test with continuity correction

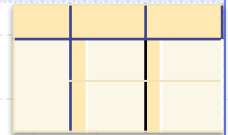
data: d7\$testA and d7\$testB

McNemar's chi-squared = 4.3478, df = 1, p-value = 0.03706

S+

```
table(d7$testA,d7$testB)
```

	no	yes
no	97	6
yes	17	80



Practical on simple analytical statistics

Using same data, test if there is a significant difference between the “weigh tafter” and the “weight before” with a paired t test.

```
t.test(d7$weightafter,d7$weightbefore,paired=TRUE)
```

Paired t-test

data: d7\$weightafter and d7\$weightbefore

t = 5.8572, df = 199, p-value = 1.922e-08

S+++

alternative hypothesis: true difference in means is not equal to 0

95 percent confidence interval:

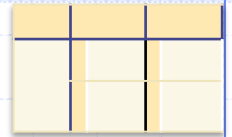
1.711051 3.447949

sample estimates:

mean of the differences

2.5795

Increasing of weight



Practical on simple analytical statistics

Using same data, test if there is a significant correlation between the “weightbefore” and the “size” with the Spearman rank correlation coefficient.

```
cor.test(d7$size,d7$weightbefore,method="spearman")
```

Spearman's rank correlation rho

data: d7\$size and d7\$weightbefore

S = 644607, p-value = **4.907e-15** **S+++**

alternative hypothesis: true rho is not equal to 0

sample estimates:

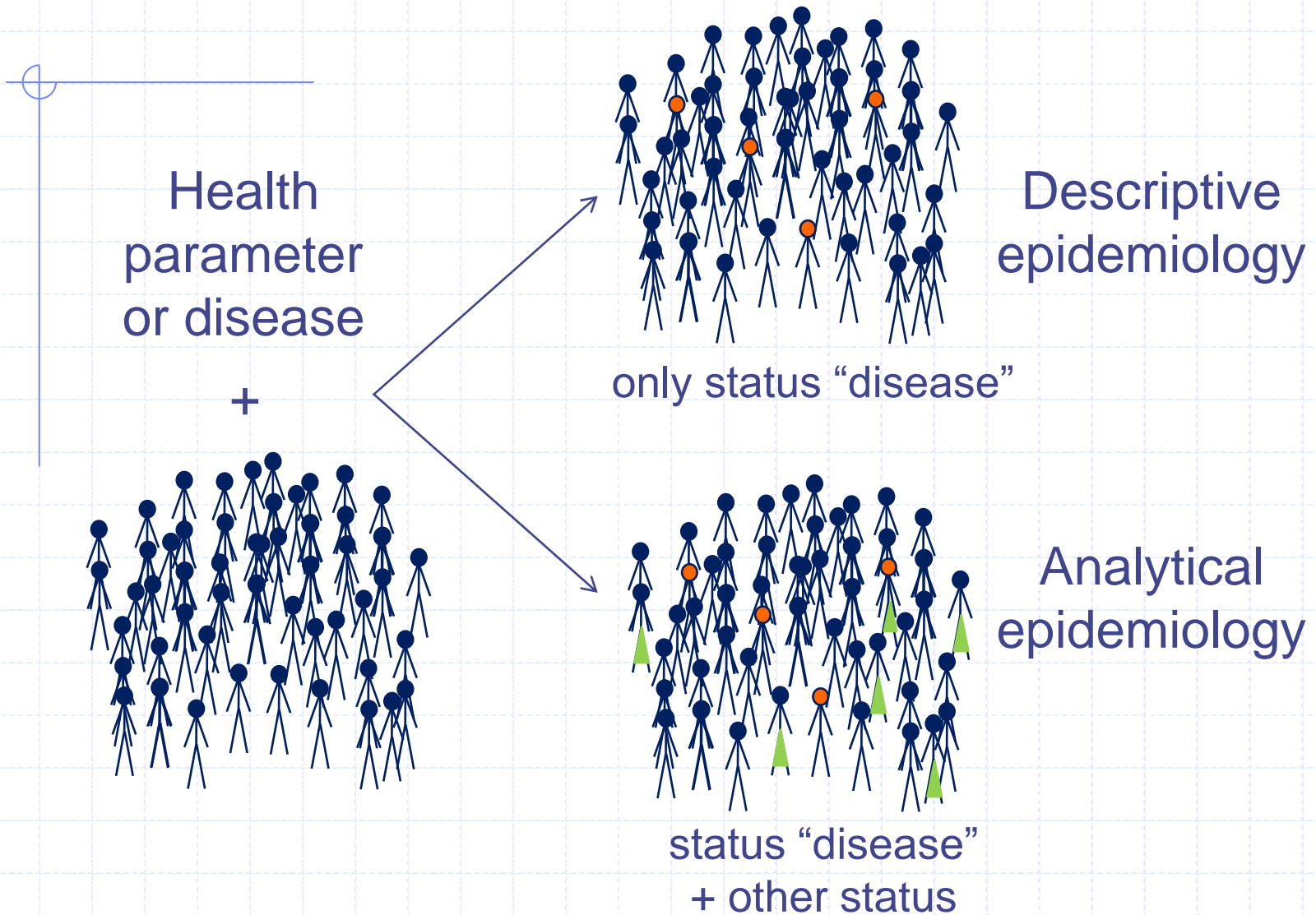
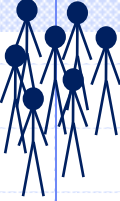
rho
0.5165327 **Positive correlation**

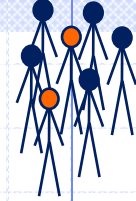
Plan

1. Principle of R language
2. Importing data
3. Simple descriptive statistics -> *Graphics*
4. Simple analytical statistics -> *tests of comparison mean and frequencies*
5. Specific tools for epidemiology
6. Specific tools for clinical study

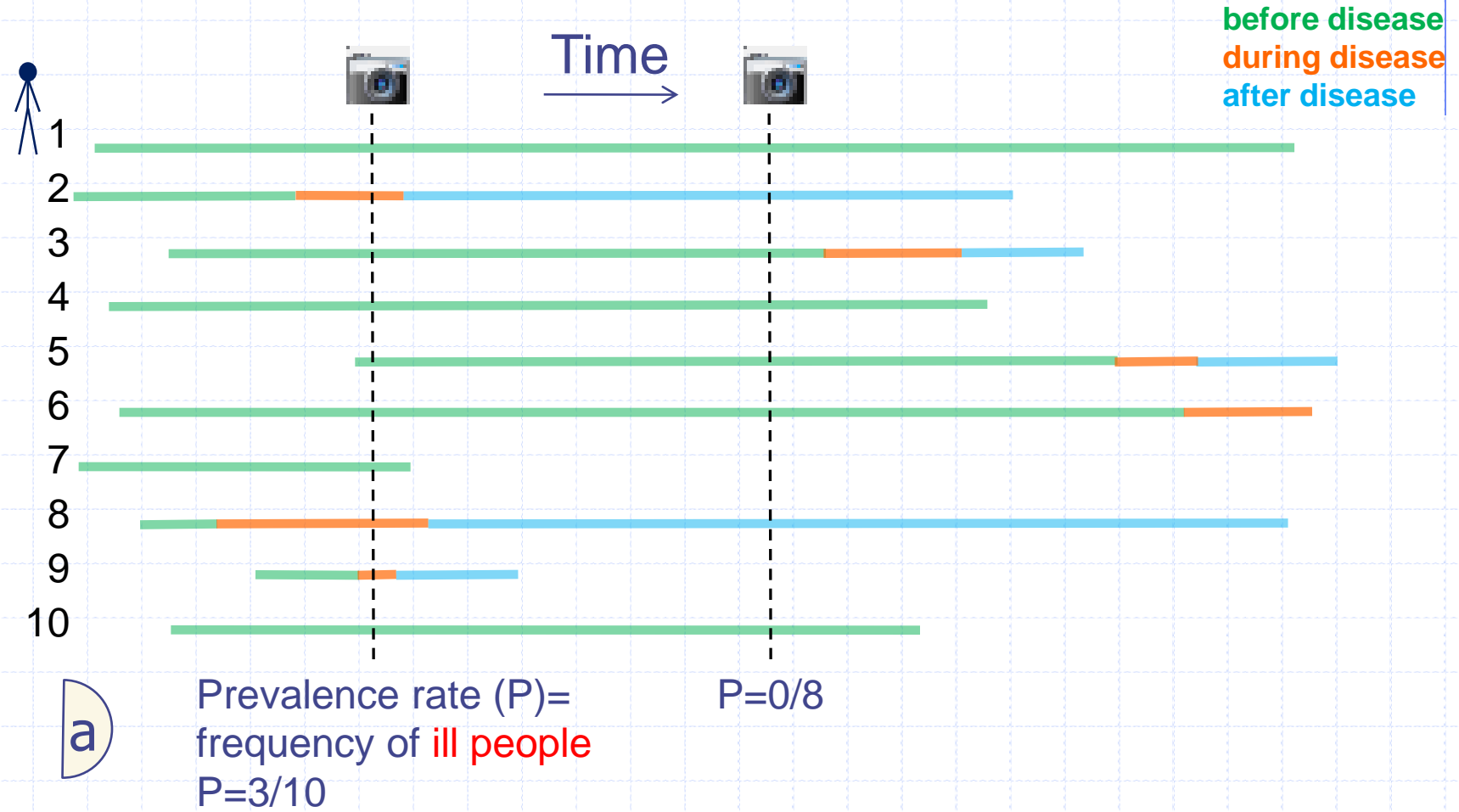


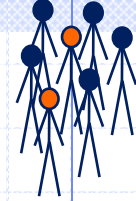
Introduction for specific tools for epidemiology





Introduction to descriptive epidemiology





Prevalence rate (P) and its Confidence Interval

`binom.test()`

Number of disease case

Total number of person

`binom.test(6,122)`

Exact binomial test

data: 6 and 122

number of successes = 6, number of trials = 122, p-value < 2.2e-16

alternative hypothesis: true probability of success is not equal to 0.5

95 percent confidence interval:

0.01825960 0.10397262

sample estimates:

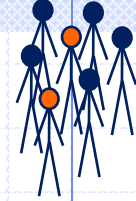
probability of success

0.04918033

Confidence Interval of
Prevalence rate P


Prevalence rate P

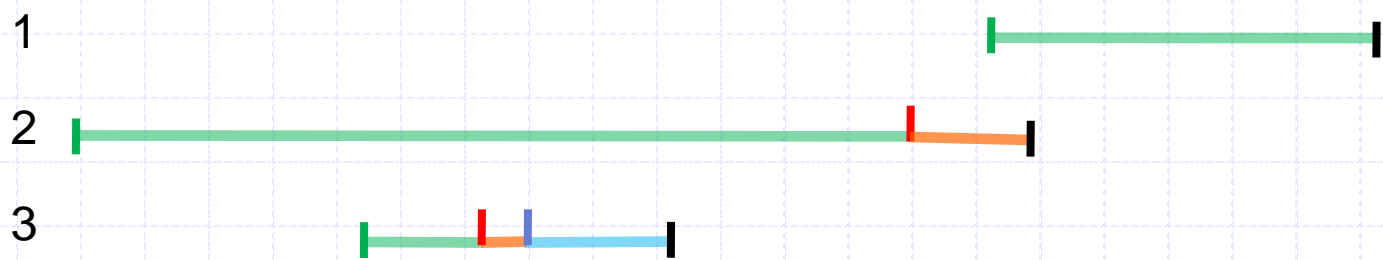
Particularly adapted
for small samples

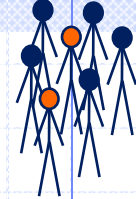


Prevalence rate (P) in cohort data -> What is it ?

```
d8[1:3,]
```

	DO 	DLN 	DI 	DEI 
1	27/10/2007	31/01/2008	<NA>	<NA>
2	23/10/2006	15/11/2007	07/10/2007	15/11/2007
3	12/01/2007	15/06/2007	24/02/2007	28/02/2007

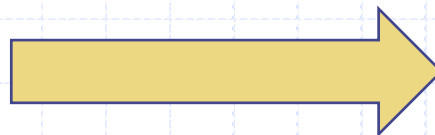




Prevalence rate (P)

in cohort data : the problem of the date

Date
22/06/1965
20/03/1968
30/08/1965
27/09/1965
22/07/1966
05/05/1964
16/01/1964
11/02/1968

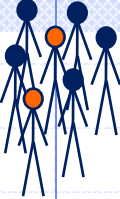


Date n
1965.472
1968.215
1965.661
1965.737
1966.553
1964.341
1964.040
1968.111

The dates are text for R so this a qualitative data. It need to transform to a quantitative value

Prevalence rate (P)

in cohort data: transform dates to fractional numbers (quantitative variables)



```
As.Date("03/07/2007", "%d/%m/%Y")  
[1] "2007-07-03"
```




Prevalence rate (P)

in cohort data : adapt the data frame
define the date of P (DPn)

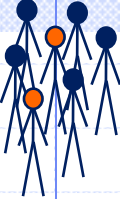
```
d8$DLNn<-as.Date(d8$DLN,"%d/%m/%Y")
d8$DIn<-as.Date(d8$DI,"%d/%m/%Y")
d8$DEIn<-as.Date(d8$DEI,"%d/%m/%Y")
d8$DOn<-as.Date(d8$DO,"%d/%m/%Y")
```

d8[1:3,]

	DO	DLN	DI	DEI	DLNn	DIn	DEIn	DOn
1	27/10/2007	31/01/2008	<NA>	<NA>	2008-01-31	<NA>	<NA>	2007-10-27
2	23/10/2006	15/11/2007	07/10/2007	15/11/2007	2007-11-15	2007-10-07	2007-11-15	2006-10-23
3	12/01/2007	15/06/2007	24/02/2007	28/02/2007	2007-06-15	2007-02-24	2007-02-28	2007-01-12

```
DPn<-as.Date("01/04/2007", "%d/%m/%Y")
```

Prevalence rate (P)



in cohort data : count the total number of persons present at the date of P (DPn)

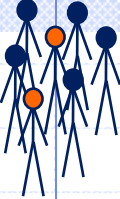
```
pres<-subset(d, DOn<=DPn & DLNn>=DPn)
npres<-nrow(pres)
npres
[1] 7
```

`subset()` return subsets of data frames which meet conditions.

The conditions here are the date of origin (DOn) is before the date for which Prevalence is calculated (DPn) and the date of the last news (DLNn) is after the DPn.

⇒ Present at the date DPn

`nrow()` return the number of rows.



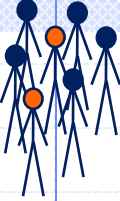
Prevalence rate (P)

in cohort data : count the total number of cases

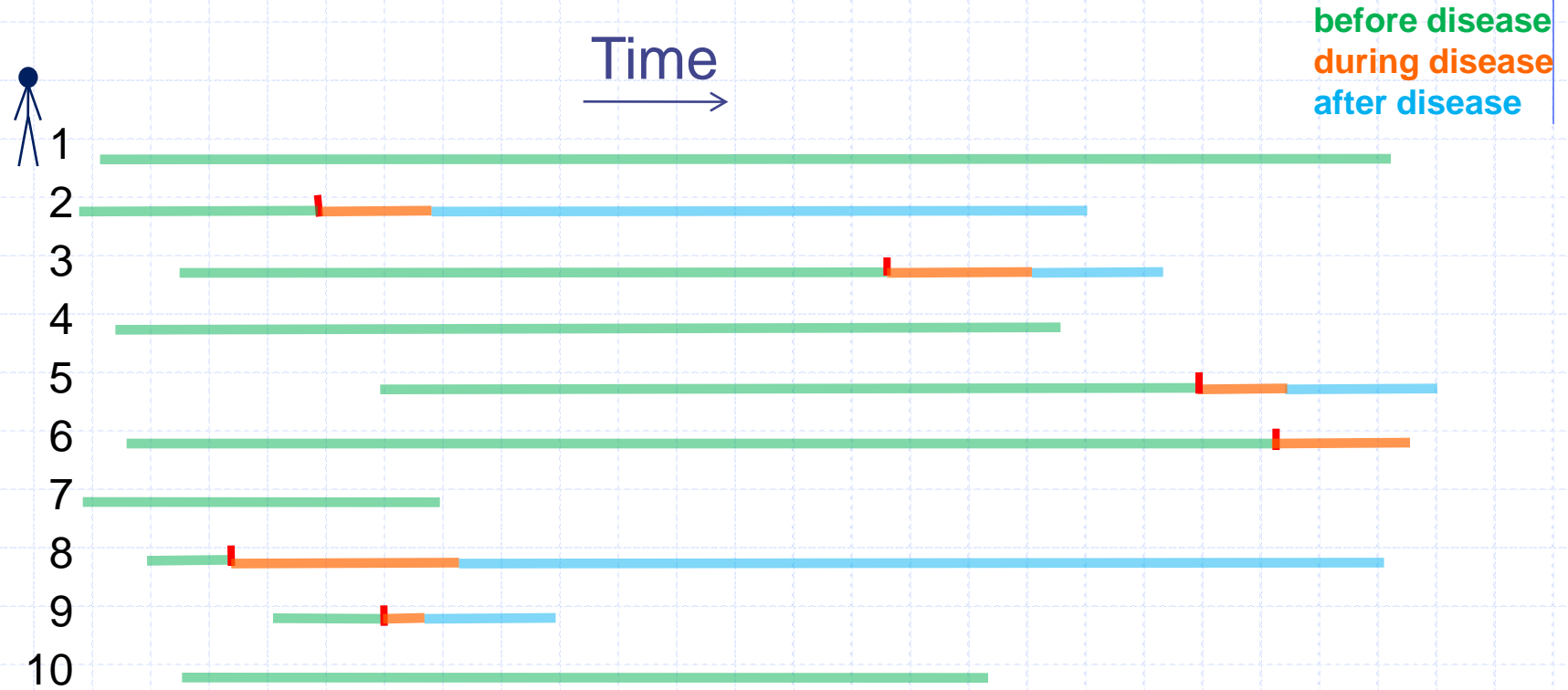
```
cases<-subset(d, DIn<=DPn & DEIn>=DPn)
ncases<-nrow(cases)
ncases
[1] 2
```

The conditions here are the date of illness (DIn) is before the date for which Prevalence is calculated (DPn) and the date of the end of illness (DEIn) is after the DPn.

⇒ Cases at the date DPn



Introduction to descriptive epidemiology

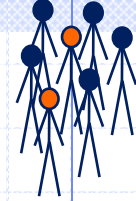


b

$$\text{Incidence rate (I)} = \frac{\text{number of new case by person-time}}{\text{sum of (—)}}$$

number of (I)
sum of (—)





Incidence rate (I) and its Confidence Interval

`pois.exact()` of the package `epitools`

The package `epitools` need to be previously download on the computer and declare it.

Number of new cases

Total number of person-time

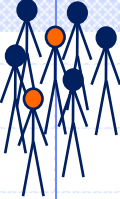
```
library(epitools)
pois.exact(3,100)
```

x	pt	rate	lower	upper	conf.level
1	3	100	0.03	0.006186712	0.08767277

Particularly adapted
for small samples

95% confidence interval of
Incidence rate I

Incidence rate I

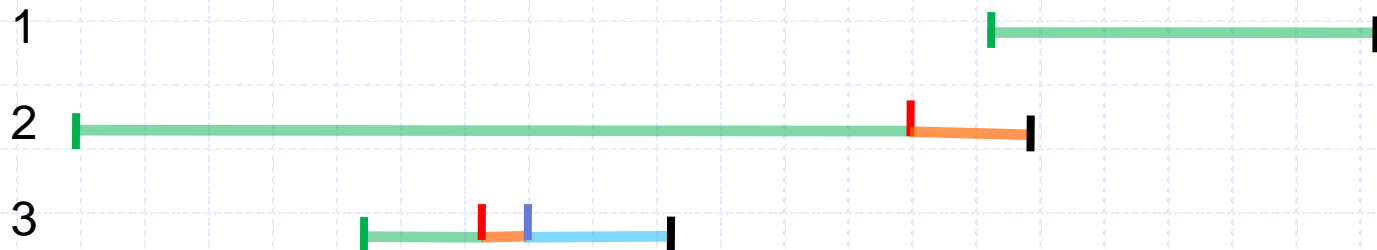


Incidence rate (I) in cohort data



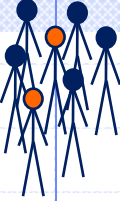
d8[1:3,]

	DO	DLN	DI	DEI
1	27/10/2007	31/01/2008	<NA>	<NA>
2	23/10/2006	15/11/2007	07/10/2007	15/11/2007
3	12/01/2007	15/06/2007	24/02/2007	28/02/2007



Participation Period (PP)

Sum of every participation time
= number of person-time (nPT)



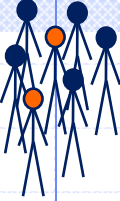
Incidence rate (I)

in cohort data : adapt the data frame

```
d8$DLNn<-as.Date(d8$DLN,"%d/%m/%Y")
d8$DIn<-as.Date(d8$DI,"%d/%m/%Y")
d8$DOn<-as.Date(d8$DO,"%d/%m/%Y")
```

```
d[1:3,]
```

	DO	DLN	DI	DLNn	DIn	DOn
1	27/10/2007	31/01/2008	<NA>	2008-01-31	<NA>	2007-10-27
2	23/10/2006	15/11/2007	07/10/2007	2007-11-15	2007-10-07	2006-10-23
3	12/01/2007	15/06/2007	24/02/2007	2007-06-15	2007-02-24	2007-01-12



Incidence rate (I) in cohort data : count the total number of person-time

```
dd<-subset(d, (DIn>DOn|is.na(DIn)))
DEn<-pmin(dd$DLNn, dd$DIn, na.rm=T)
DBn<-dd$DOn
PP<-DEn-DBn
nPT<-sum(PP)
nPT
[1] 6.551677
```

Na.rm=T removes all the NA values

The conditions here are the date of illness(DIn) is after the date of origin (DOn) or the date of illness is not available (person still healthy).


The date of the end of time of participation (DEn) is minimum between DLNn et DIn.

The date of the beginning of time of participation (DBn) is DOn

The participation periods (PP) is the difference between DEn and DBn

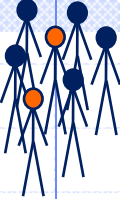
The number of Person-Time (nPT) is the sum of every participation period (PP)



Incidence rate (I) 
in cohort data : count the total
number of cases

```
newcases<-subset(dd,DIn>DOn)  
nnewcases<-nrow(nnewcases)  
nnewcases  
[1] 4
```

The conditions here are the date of illness (DIn) is after
the date of Origin (DOn).
⇒ New cases during the study



Practical on simple descriptive epidemiology

In a survey on 1538 person, 19 have antibodies against a disease. What are the seroprevalence rate of the disease and its 95% confidence interval?



```
binom.test(19,1538)
```

Exact binomial test

data: 19 and 1538

number of successes = 19, number of trials = 1538, p-value < 2.2e-16

alternative hypothesis: true probability of success is not equal to 0.5

95 percent confidence interval:

0.007453676 0.019224878

sample estimates:

probability of success

0.01235371

P=1.24% [0.75%,1.92%]



Practical on simple descriptive epidemiology

In the same survey, what is the 90% confidence interval of the sero-prevalence rate? Use the argument `conf.level=0.9`.



```
binom.test(19,1538,conf.level=0.9)
```

Exact binomial test

```
data: 19 and 1538
```

```
number of successes = 19, number of trials = 1538, p-value < 2.2e-16
```

```
alternative hypothesis: true probability of success is not equal to 0.5
```

```
90 percent confidence interval:
```

```
0.008104379 0.018074569
```

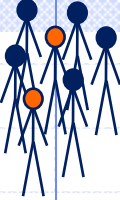
More narrow confident interval

```
sample estimates:
```

```
probability of success
```

```
0.01235371
```

P=1.24% [0.81%,1.81%]



Practical on simple descriptive epidemiology

In a cohort study, what are the prevalence rate and its 95% confidence interval at the 1st April 2007? (“cohort.txt”)



```
d8$DLNn<-as.Date(d8$DLN,"%d/%m/%Y")
d8$DIn<-as.Date(d8$DI,"%d/%m/%Y")
d8$DEIn<-as.Date(d8$DEI,"%d/%m/%Y")
d8$DOn<-as.Date(d8$DO,"%d/%m/%Y")
DPn<-as.Date("01/04/2007", "%d/%m/%Y")
pres<-subset(d8, DOn<=DPn & DLNn>=DPn)
(npres<-nrow(pres))
case<-subset(d8, DIn<=DPn & DEIn>=DPn)
(ncase<-nrow(case))
binom.test(ncase,npres)
```



Practical on simple descriptive epidemiology

In a cohort study, what are the prevalence rate and its 95% confidence interval at the 1st April 2007? (“cohort.txt”)

```
binom.test(ncase,npres)
```

```
Exact binomial test
```

```
data: ncase and npres
```

```
number of successes = 2, number of trials = 7 p-value = 0.4531
```

```
alternative hypothesis: true probability of success is not  
equal to 0.5
```

```
95 percent confidence interval:
```

```
0.03669257 0.70957914
```

```
sample estimates:
```

```
probability of success
```

```
0.2857143
```

P = 29% [3.7%,71%]



Practical on simple descriptive epidemiology

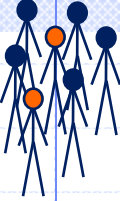
In a cohort study, the total number of person-time at risk is 34590 women-month, a specific pathology is observed 15. What are the incidence rate and its 95% confidence interval?



```
library(epitools)
pois.exact(15,34590)
```

	x	pt	rate	lower	upper	conf.level
1	15	34590	0.0004336513	0.0002427113	0.000715242	0.95

I=433 [243,715] cases for 100,000 women-month

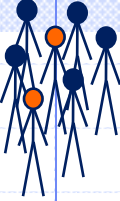


Practical on simple descriptive epidemiology

In the cohort study of exercise 6e, what are the incidence rate and its 95% CI during all the study ? (“cohort.txt”)



```
dd<-subset(d,DIn>DOn|is.na(DIn) )  
DEn<-pmin(dd$DLNn,dd$DIn,na.rm=T)  
DBn<-dd$DOn  
PP<-DEn-DBn  
nPT<-sum(PP)  
newcases<-subset(dd,DIn>DOn)  
nnewcases<-nrow(newcases)  
pois.exact(nnewcases,nPT)
```



Practical on simple descriptive epidemiology

In the cohort study of exercise 6e, what are the incidence rate in person-year and its 95% CI during all the study? (“cohort.txt”)

```
pois.exact(nnewcases,nPT)
```

x	pt	rate	lower	upper	conf.level
8	1315	0.00608365	0.002626488	0.01198722	0.95

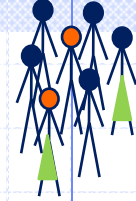
```
pois.exact(nnewcases,nPT/365.25)
```

x	pt	rate	lower	upper	conf.level
1	8	3.600274	2.222053	0.9593247	4.378331

I=2.2 [0.96,4.4] new cases for 1 person-years



The unit of number of Person-time is the person-days because the unit of the time is the day (Cf. as.Date)



Introduction to analytical epidemiology

-> association between exposition and disease

Ex. : Lung cancer

Ex. : tabacco

		Disease		
		-	+	
Exposition	-	a	b	n_0
	+	c	d	n_1
		m_0	m_1	

Cohort survey

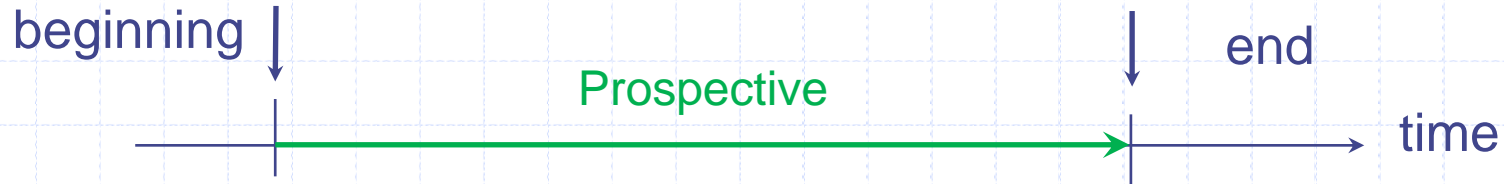
a) b)

Case-control study

c)



Cohort survey and Risk Ratio (RR)



$$\begin{aligned}
 n_1 = E_+ D_- & \rightarrow E_+ D_- = c \\
 & \rightarrow E_+ D_+ = d \\
 n_0 = E_- D_- & \rightarrow E_- D_- = a \\
 & \rightarrow E_- D_+ = b
 \end{aligned}$$

		Disease		
		-	+	
Exposition	-	a	b	n_0
	+	c	d	n_1
		m_0	m_1	

$$\text{Risk to get ill in exposed} = R_1 = \frac{d}{n_1}$$

$$\text{Risk to be ill in not-exposed} = R_0 = \frac{b}{n_0}$$

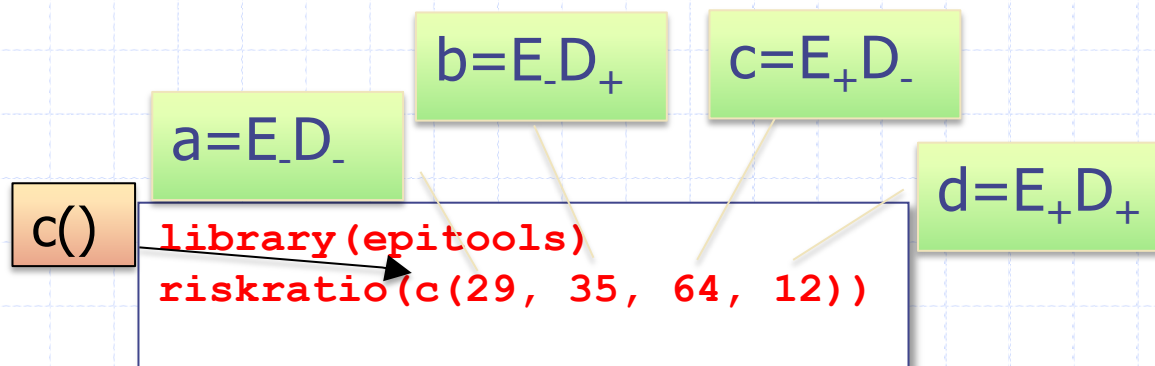
$$\text{Risk Ratio} = RR = \frac{R_1}{R_0}$$



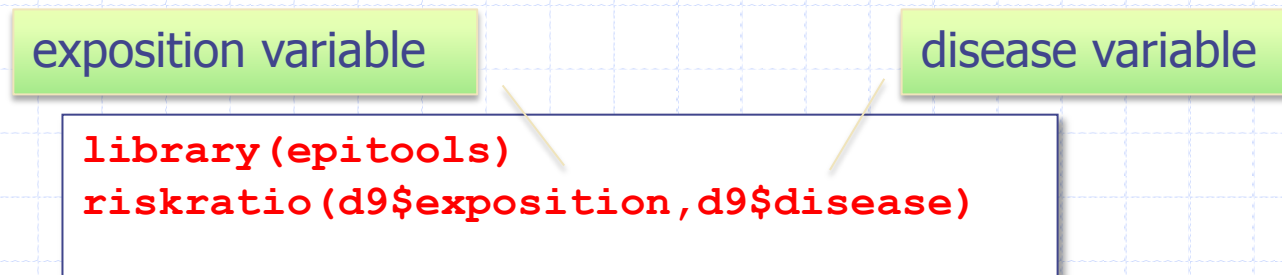
Risk Ratio (RR) and its Confidence Interval

`riskratio()` of the package `epitools`

The package `epitools` need to be previously download on the computer and declare it.



Or





Risk Ratio (RR) and its Confidence Interval

```
library(epitools)
riskratio(c(29, 35, 64, 12))
```

```
$data
```

Predictor	Outcome	Disease1	Disease2	Total
Exposed1		29	35	64
Exposed2		64	12	76
Total		93	47	140

Risk Ratio RR

```
$measure
```

Predictor	estimate	lower	upper
Exposed1	1.0000000	NA	NA
Exposed2	0.2887218	0.1640857	0.5080288

95% confidence interval of
Risk Ratio RR

```
$p.value
```

Predictor	midp.exact	fisher.exact	chi.square
Exposed1	NA	NA	NA
Exposed2	1.282338e-06	1.876171e-06	1.203275e-06

p value of fisher exact test
-> RR≠1?

```
$correction
[1] FALSE
```

```
attr(,"method")
```

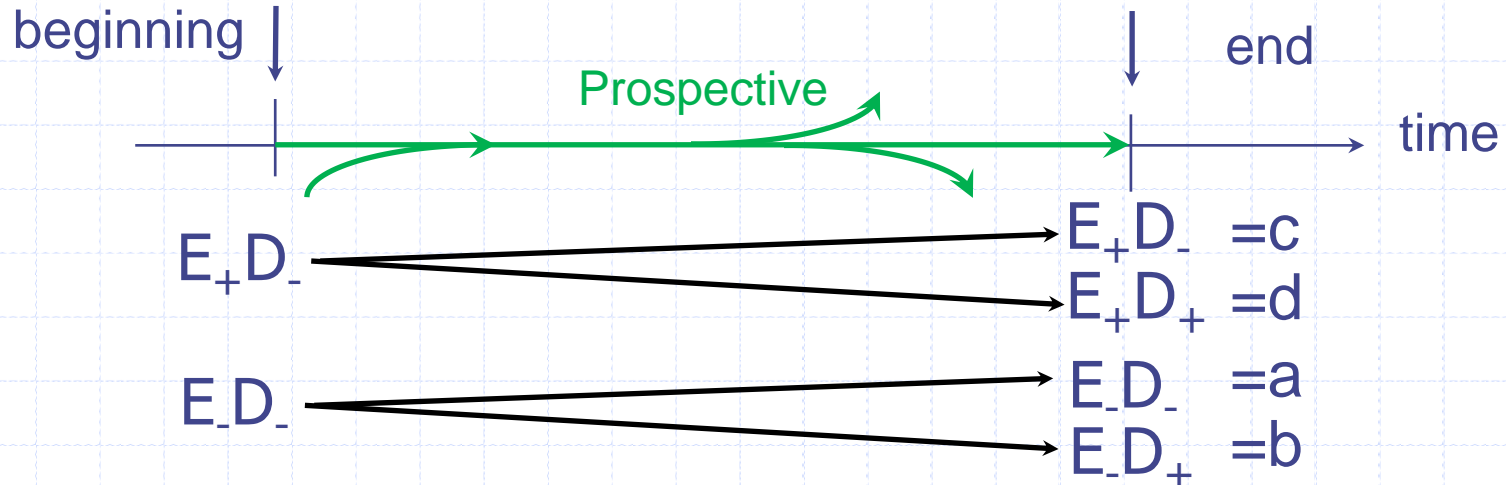
```
[1] "Unconditional MLE & normal approximation (Wald) CI"
```



If small sample



Cohort survey and Incidence Rate Ratio (IRR)



N_{pt1} = number of person-time in exposed

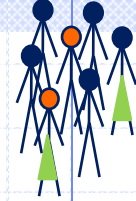
I_1 = Incidence rate in exposed = a/N_{pt1}

N_{pt0} = number of person-time in not-exposed

I_0 = Incidence rate in not-exposed = b/N_{pt0}

~~$$RR = \frac{R_1}{R_0}$$~~

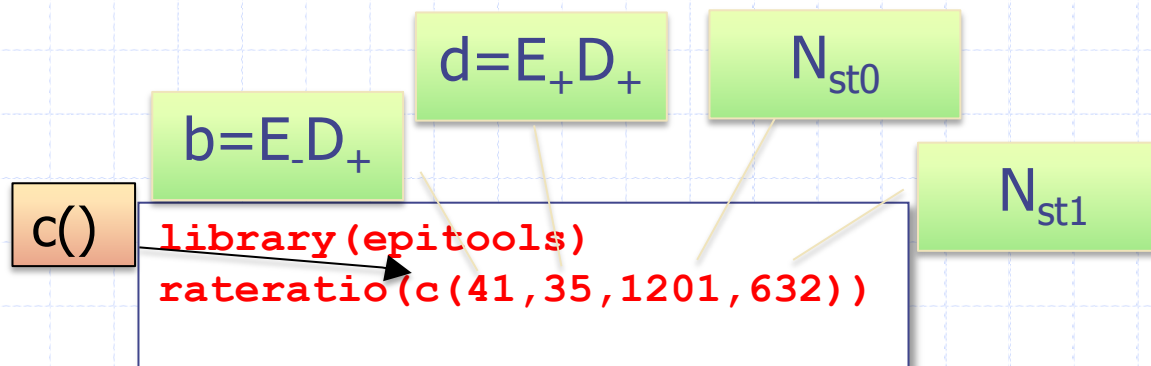
$$\text{Incidence Rate Ratio (IRR)} = \frac{I_1}{I_0} = \frac{\frac{a}{N_{pt1}}}{\frac{b}{N_{pt0}}}$$



Incidence Rate Ratio (IRR) and its Confidence Interval

`rateratio()` of the package `epitools`

The package `epitools` need to be previously download on the computer and declare it.





Incidence Rate Ratio (IRR) and its Confidence Interval

```
library(epitools)
rateratio(c(41,35,1201,632))
```

```
$data
```

Predictor	Outcome	
	Cases	Person-time
Exposed1	41	1201
Exposed2	35	632
Total	76	1833

```
$measure
```

```
rate ratio with 95% C.I.
```

Predictor	estimate	lower	upper
Exposed1	1.000000	NA	NA
Exposed2	1.623489	1.027673	2.550396

```
$p.value
```

```
two-sided
```

Predictor	midp.exact	wald
Exposed1	NA	NA
Exposed2	0.03797212	0.03377064

```
attr(,"method")
```

```
[1] "Median unbiased estimate & mid-p exact CI"
```

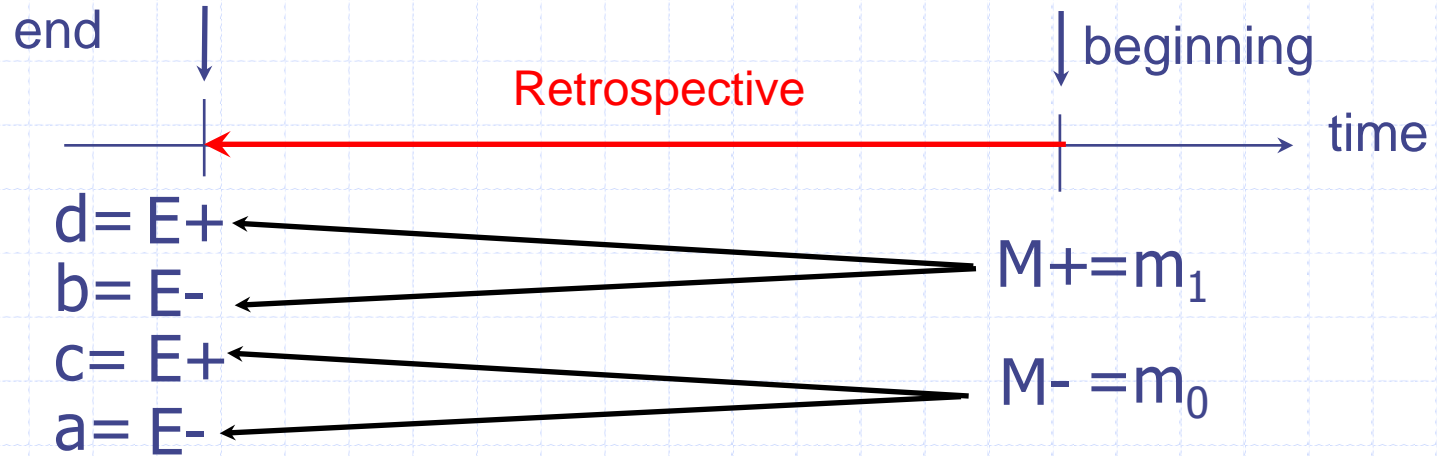
Incidence Rate Ratio IRR

95% confidence interval of
Incidence Rate Ratio RR

p value of midp exact test
-> IRR≠1?



Case-control Study and Odds Ratio (OR)



		Disease		
		-	+	
Exposition	-	a	b	n ₀
	+	c	d	n ₁
		m ₀	m ₁	

~~$$RR = \frac{R_1}{R_0}$$~~

$$\text{Odds Ratio} = OR = \frac{ad}{bc}$$



Relation between Odds Ratio (OR) and Risk Ratio (RR)

$$RR = \frac{OR}{1 + R_0(OR - 1)}$$

with R_0 risk to get the
disease in not-exposed

if $R_0 \ll 1$ then $RR \approx OR$

if $OR > 1$ then $OR > RR > 1$
Exposition linked to the disease

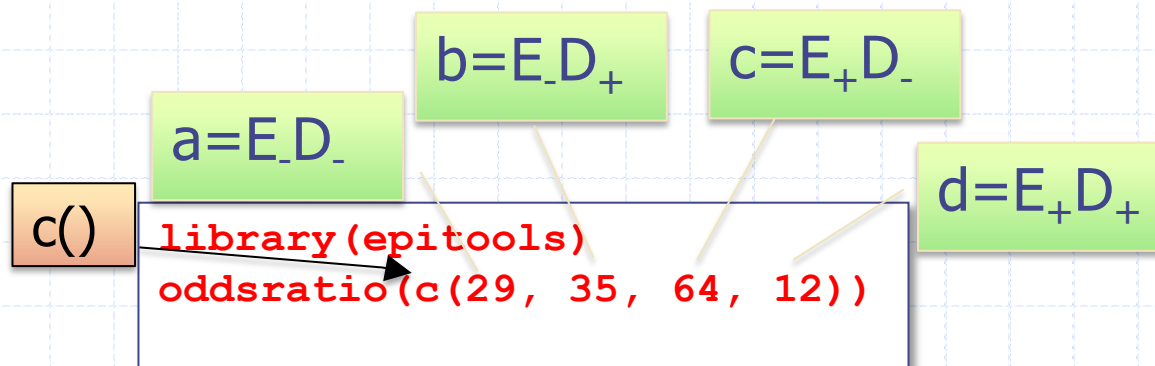
si $OR < 1$ alors $OR < RR < 1$
Exposition linked to the protection



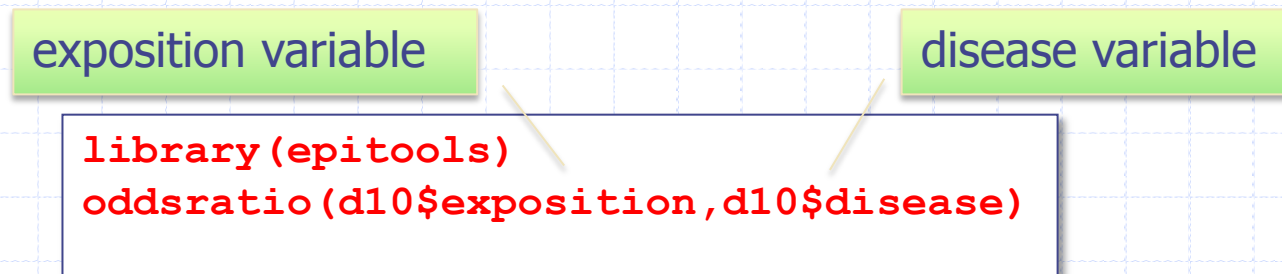
OddsRatio (OR) and its Confidence Interval

`oddsratio()` of the package `epitools`

The package `epitools` need to be previously download on the computer and declare it.



Or





Odds Ratio (OR) and its Confidence Interval

```
library(epitools)
oddsratio(c(29, 35, 64, 12))
```

```
$data
```

Predictor	Outcome		Total
	Disease1	Disease2	
Exposed1	29	35	64
Exposed2	64	12	76
Total	93	47	140

Odds Ratio OR

```
$measure
```

Predictor	odds ratio with 95% C.I.		
	estimate	lower	upper
Exposed1	1.0000000	NA	NA
Exposed2	0.1587889	0.06939932	0.3427073

95% confidence interval of
Odds Ratio OR

```
$p.value
```

Predictor	two-sided		
	midp.exact	fisher.exact	chi.square
Exposed1	NA	NA	NA
Exposed2	1.282338e-06	1.876171e-06	1.203275e-06

p value of fisher exact test
-> OR≠1?

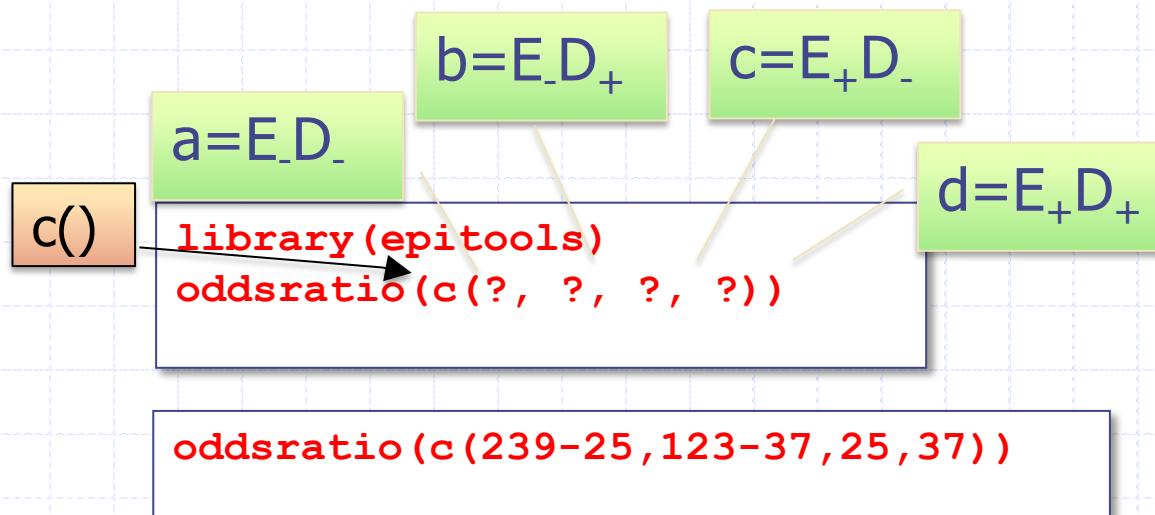
```
$correction
[1] FALSE
```

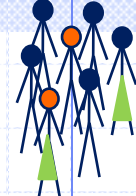
```
attr(,"method")
[1] "median-unbiased estimate & mid-p exact CI"
```



Practical on simple analytical epidemiology

In a case-control study, among 123 cases, 37 are smokers and among 239 controls, 25 are smokers. What are the Odds Ratio and its 95% Confidence Interval?





Practical on simple analytical epidemiology

```
$data
Outcome
Predictor Disease1 Disease2 Total
Exposed1    214      86    300
Exposed2     25      37     62
Total       239    123    362
```

```
$measure
odds ratio with 95% C.I.
Predictor estimate lower upper
Exposed1 1.000000 NA NA
Exposed2 3.660792 2.085971 6.525364
```

OR=3.66 [2.09,6.52]

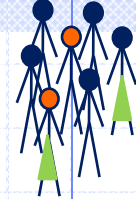
```
$p.value
two-sided
Predictor midp.exact fisher.exact chi.square
Exposed1 NA NA NA
Exposed2 5.765939e-06 6.427031e-06 2.689288e-06
```

OR≠1 S+++
(p<0.001)

```
$correction
[1] FALSE
```

```
attr(,"method")
[1] "median-unbiased estimate & mid-p exact CI"
```

Exposition is linked
with disease (OR>1)



Practical on simple analytical epidemiology

The results of a cohort study are summarized in the file “cohort2.txt”. The column exposition and disease correspond respectively to the information about the people which smoke or not and the information about the people which get ill or not. What are the Relative Risk and its 95% Confidence Interval?

exposition variable

disease variable

```
library(epitools)  
riskratio( ? , ? )
```

```
riskratio(d11$exposition,d11$disease)
```



Practical on simple analytical epidemiology

\$data

Predictor	Outcome		
	no	yes	Total
no	375	25	400
yes	351	49	400
Total	726	74	800

\$measure

risk ratio with 95% C.I.

Predictor	estimate	lower	upper
no	1.00	NA	NA
yes	1.96	1.235641	3.108994

RR=1.96 [1.23,3.11]

\$p.value

two-sided

Predictor	midp.exact	fisher.exact	chi.square
no	NA	NA	NA
yes	0.003414235	0.004725127	0.003404036

**RR≠1 S++
(p<0.01)**

\$correction

[1] FALSE

attr(,"method")

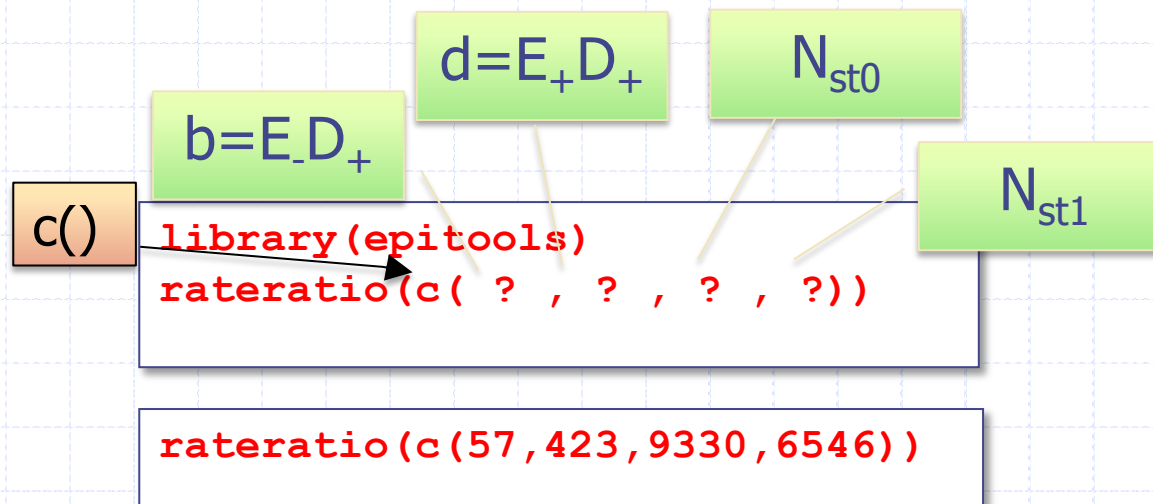
[1] "Unconditional MLE & normal approximation (Wald) CI"

**Exposition is linked
with disease (RR>1)**



Practical on simple analytical epidemiology

The results of a cohort study, all the people don't stay the same period in the cohort. In the exposed population (smoker), the total number of person-time at risk is 6546 people-years and 423 get ill. In the not-exposed population (no-smoker), the total number of person-time at risk is 9330 people-years and 57 get ill. What are the Incidence rate ratio and its 95% Confidence Interval?





Practical on simple analytical epidemiology

\$data

Predictor	Outcome	
	Cases	Person-time
Exposed1	57	9330
Exposed2	423	6546
Total	480	15876

\$measure

rate ratio with 95% C.I.

Predictor	estimate	lower	upper
Exposed1	1.00000	NA	NA
Exposed2	10.55062	8.07544	14.05593

IRR=10.1 [8.08,14.1]

\$p.value

two-sided

Predictor	midp.exact	wald
Exposed1	NA	NA
Exposed2	0	0

**IRR≠1 S+++
(p<0.001)**

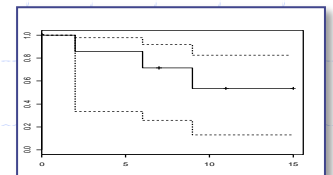
attr(,"method")

[1] "Median unbiased estimate & mid-p exact CI"

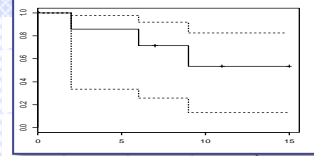
**Exposition is linked
with disease (IRR>1)**

Plan

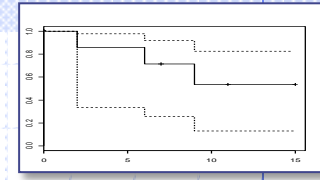
1. Principle of R language
2. Importing data
3. Simple descriptive statistics -> *Graphics*
4. Simple analytical statistics -> *tests of comparison mean and frequencies*
5. Specific tools for epidemiology
6. Specific tools for clinical study



Clinical study -> survival curves

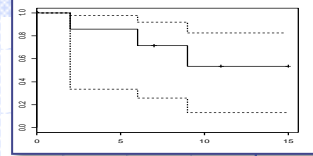


- ◆ Understand the principle of the survival curves and their confidence interval
- ◆ Know how realize survival curves from data frame with R.



Interest of survival curves

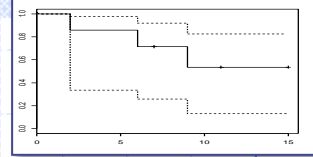
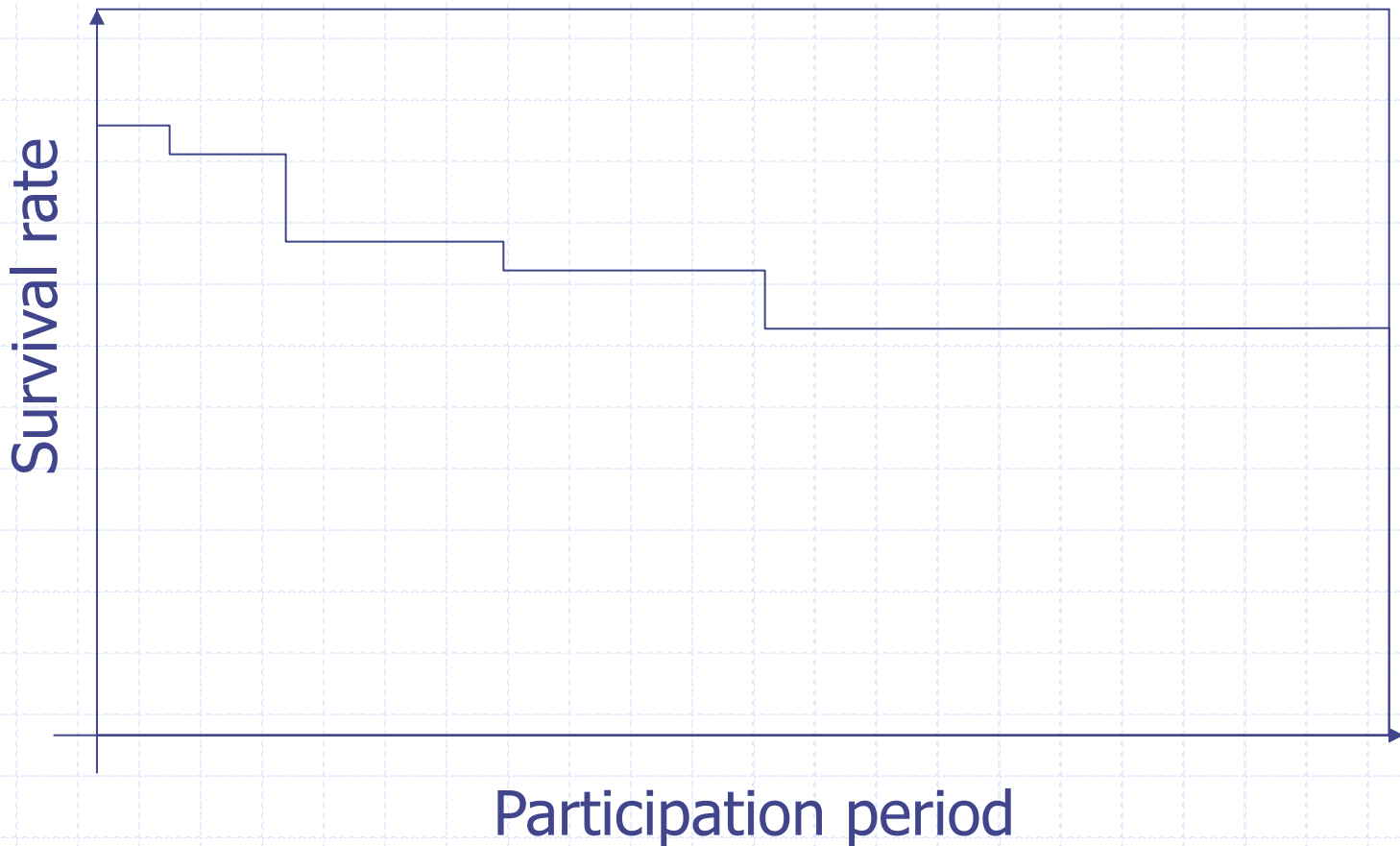
- ◆ Describe the survival rate in a sample;
- ◆ Estimate the survival rate in the population;
- ◆ Compare survival rates according groups (therapeutic and epidemiologic research).

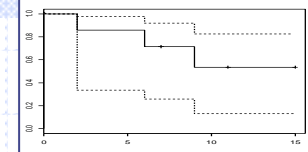


Survival data

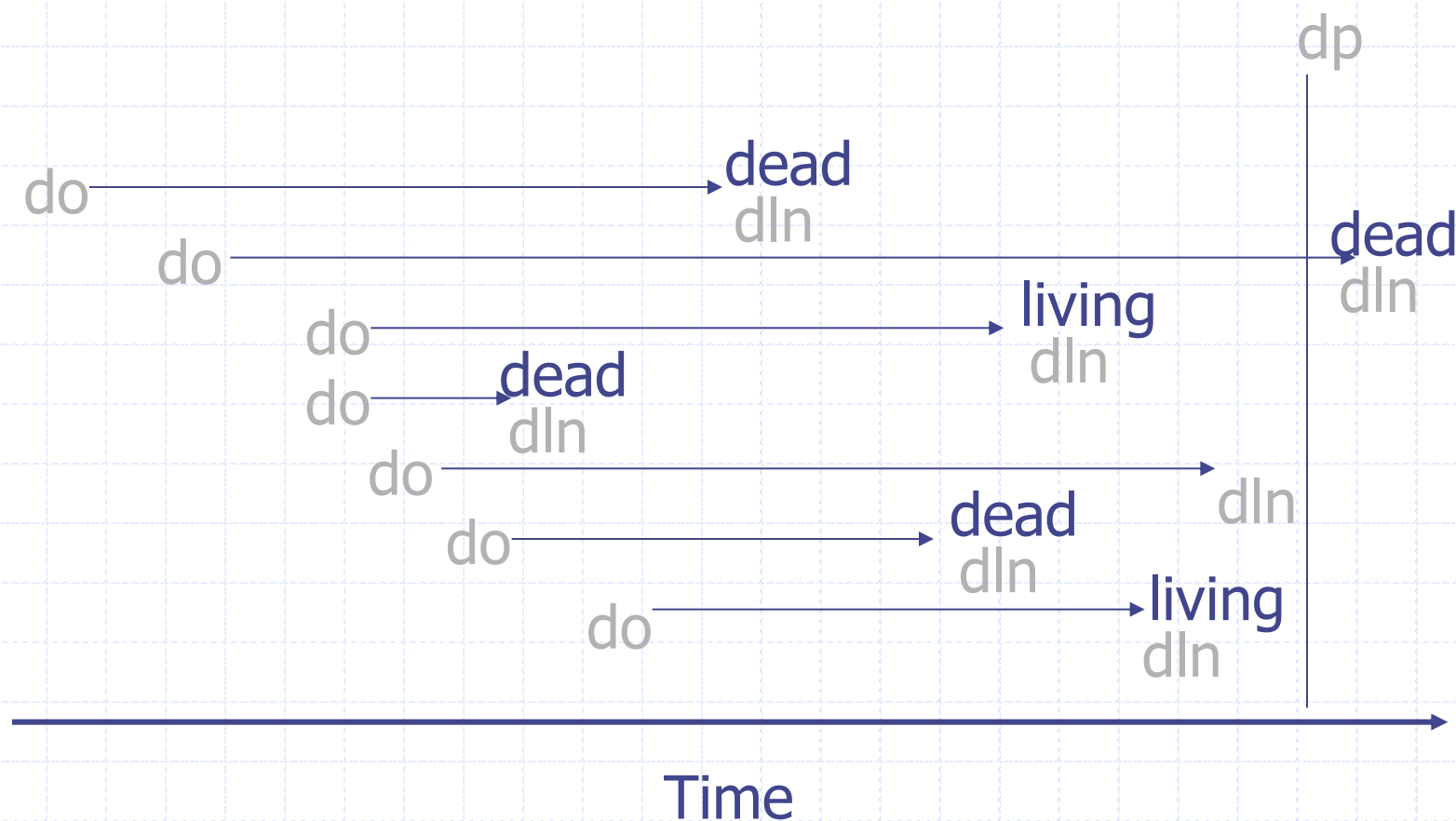
- ◆ Survival : time between two dates. The final date is not necessarily the date of death e.g., date of first relapse, of the first complication,...
- ◆ Particularity of survival data: possibility de censored data
- ◆ Survival rate : probability to be alive (or healthy)
≠ 1-mortality rate
mortality rate is an incidence rate
- ◆ Survival curve : Representation of the survival rate according the time of participation

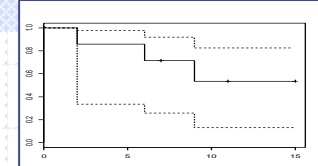
Survival curve



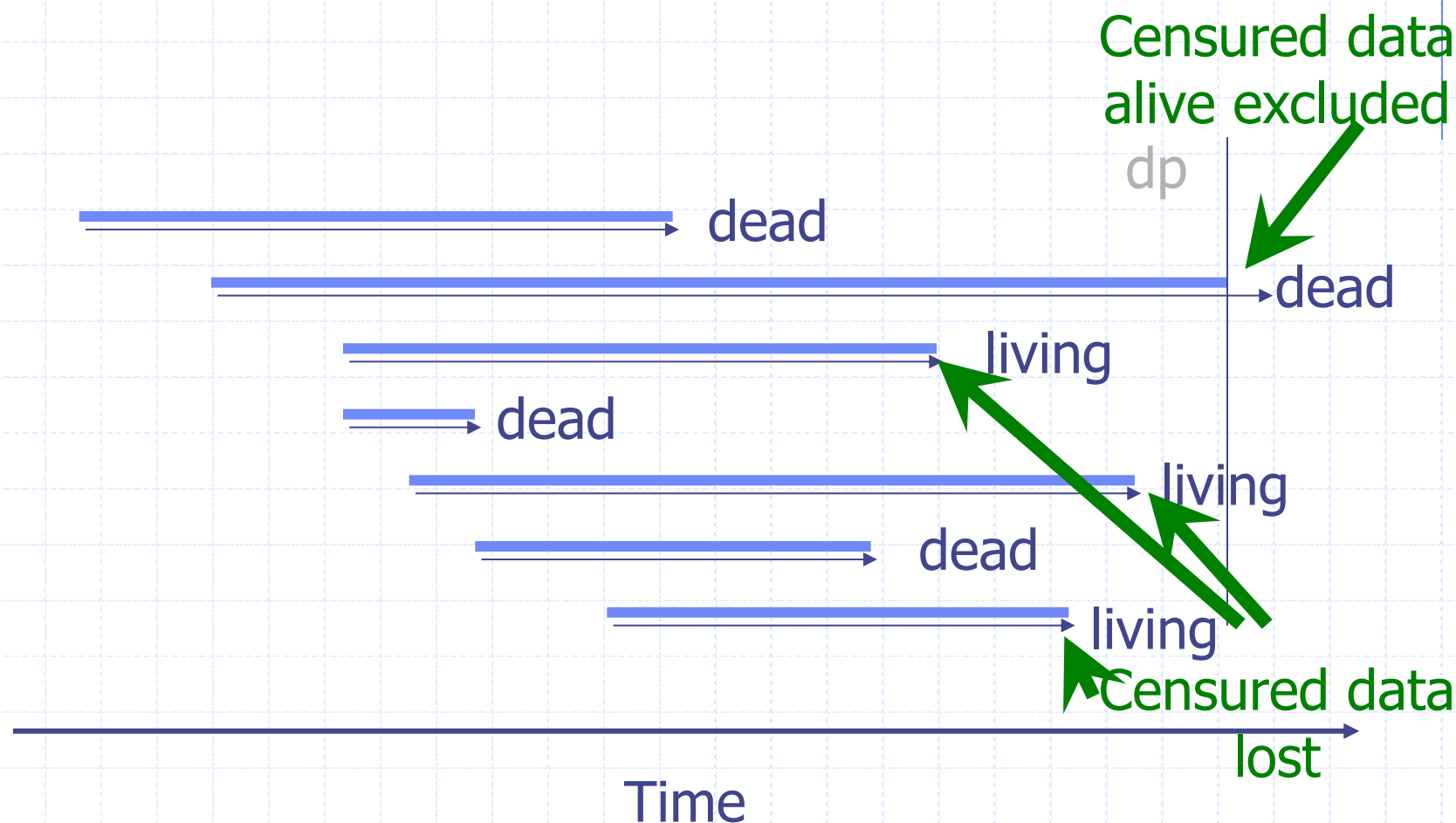


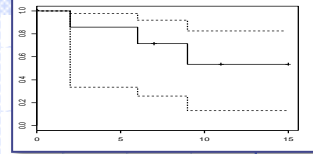
Participation period



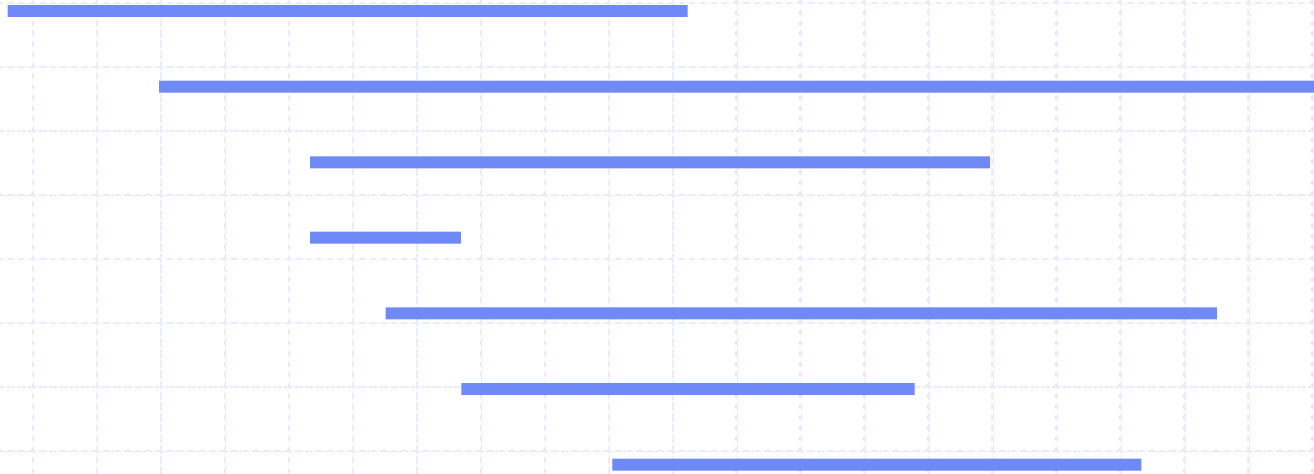


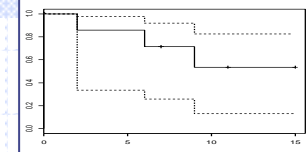
Participation period and **censure**



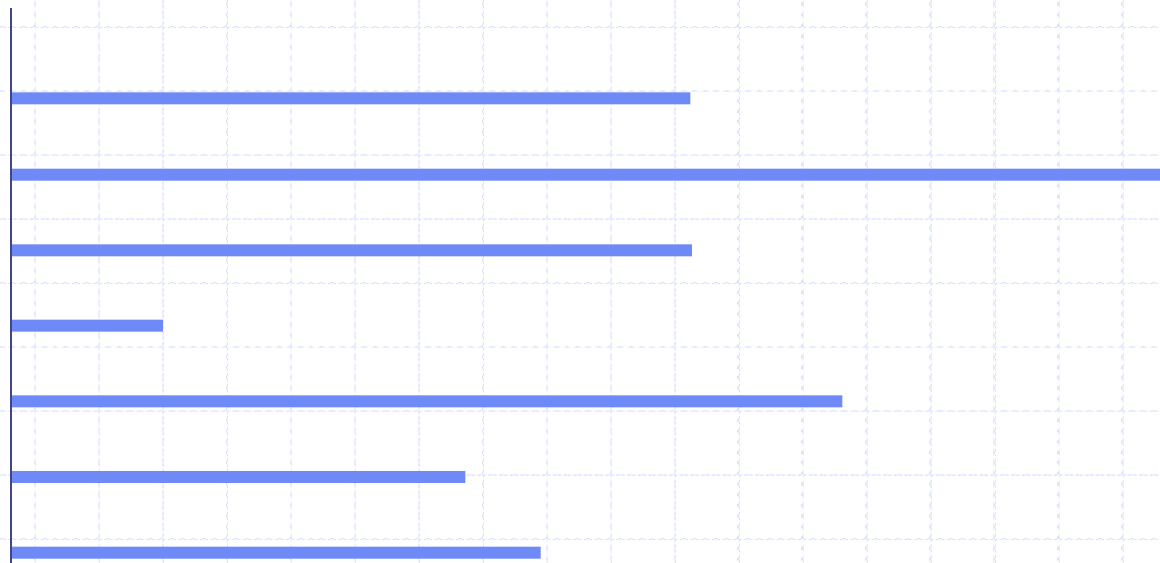


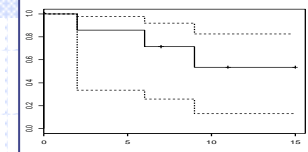
Participation period from 0



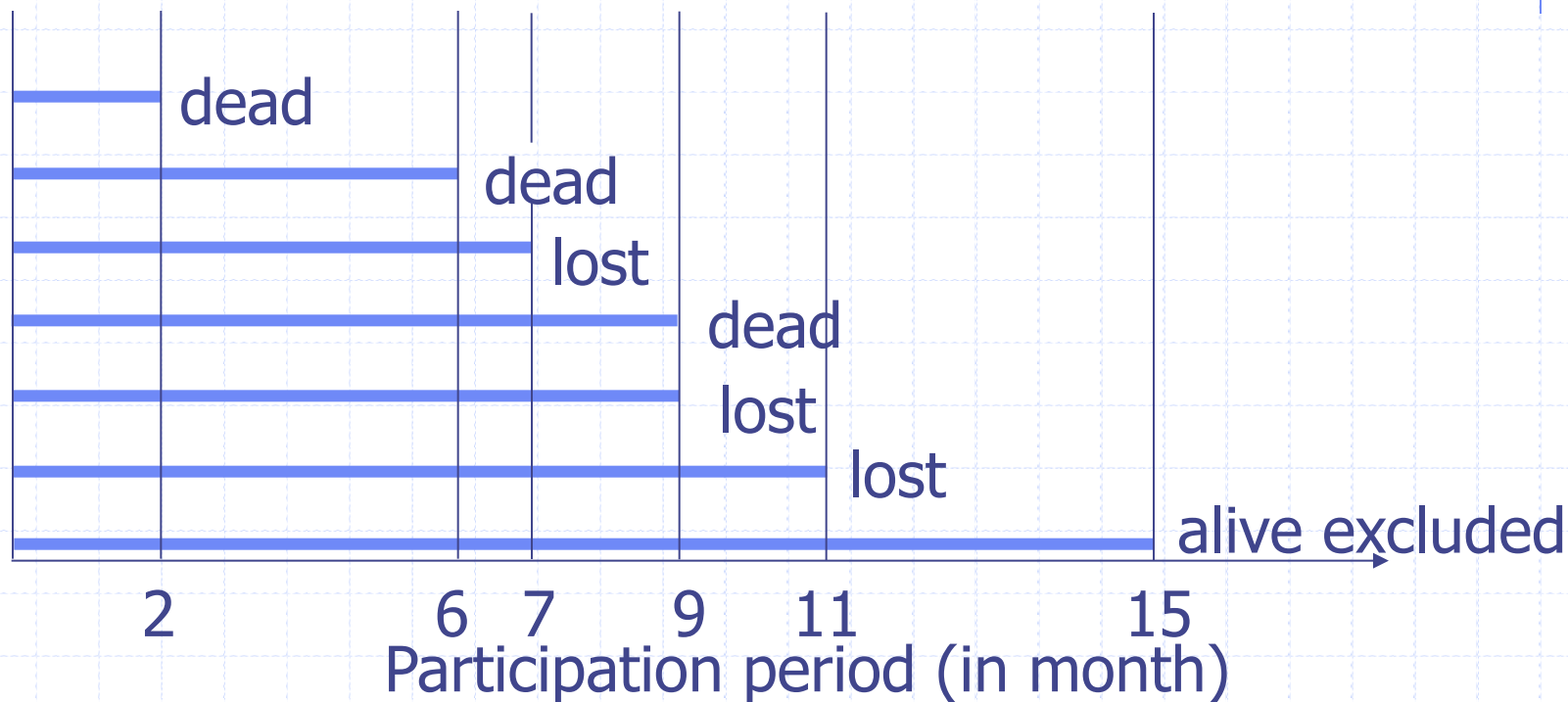


Ordered participation period

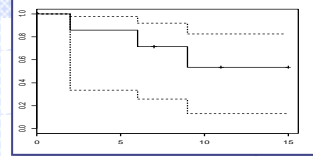




Summary of participation period



→ Survival curve



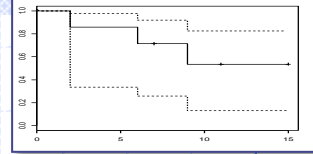
Survival rate

The survival rate represent the probability to be still alive a the time t . There are different non parametric methods to estimate the survival rate:

- ◆ Kaplan-Meier method
- ◆ Actuarial method

These methods are based on the principle of the conditional probabilities.

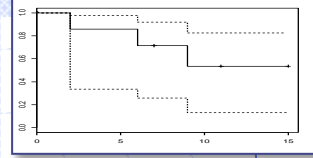
- ◆ Method of direct calculus



Kaplan-Meier method

In this method, the survival rate is calculated only at t_i when a time of participation is finishing.

- If the time of participation is finishing by a dead then the survival decrease.
- Otherwise, the survival rate is constant.

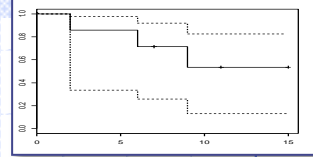


Principle of Kaplan-Meier method

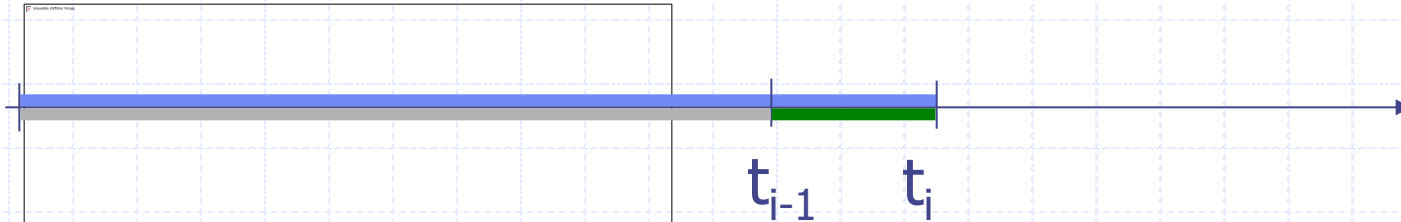
So, the probability to be alive at time t_i is equal to:

- the probability to be alive at t_{i-1} multiplied by
- the probability to be alive at t_i conditionally to be alive at t_{i-1} .

For the time between t_i and t_{i+1} excluded, the probability to be alive is constant.



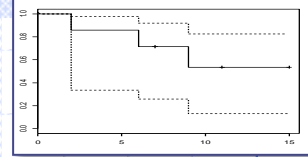
Conditionnall probabilities



$$P(S_{t_i}) = P(S_{t_{i-1}}) \times P(S_{t_i} | t_{i-1})$$

$$S_i = S_{i-1} \times S_{i|i-1}$$

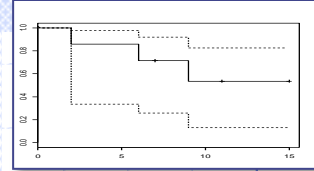
Calculus of survival rate with the Kaplan-Meier method



$$S_{i|i-1} = \frac{N_i - D_i}{N_i}$$

$$S_i = \frac{N_1 - D_1}{N_1} \times \frac{N_2 - D_2}{N_2} \times \dots \times \frac{N_i - D_i}{N_i}$$

Calculus of CI of survival rates with the Kaplan-Meier method



$$\text{IC de } S_i = S_i \pm u_{1-\alpha/2} \sqrt{\text{var } S_i}$$

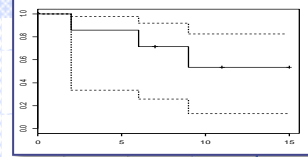
Variance of survival rate
or Greenwood variance

$$\text{var } S_i = S_i^2 \left[\frac{D_1}{N_1 (N_1 - D_1)} + \dots + \frac{D_i}{N_i (N_i - D_i)} \right]$$



Be careful to the approximation by normal law

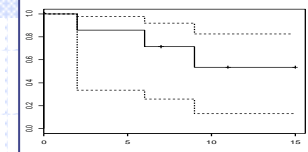
Calculus of CI of survival rates with the Kaplan-Meier method



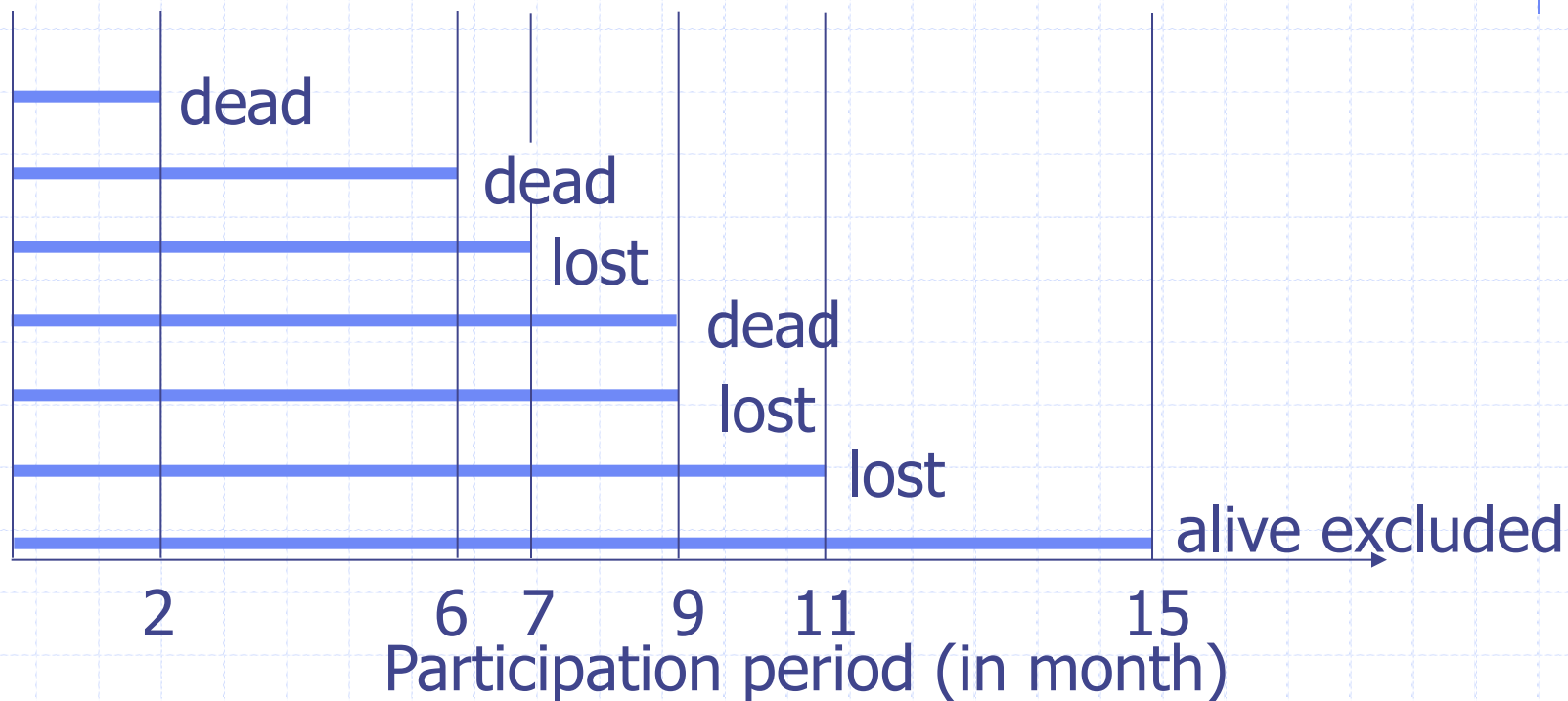
If the approximation by normal law is not proved -> correction of Rothman or `surfit()` in R

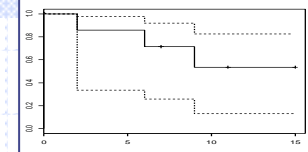
$$\text{IC de } S_i = \frac{M}{M + u_{1-\alpha/2}^2} \left[S_i + \frac{u_{1-\alpha/2}^2}{2M} \pm u_{1-\alpha/2} \sqrt{\text{var } S_i + \frac{u_{1-\alpha/2}^2}{4M^2}} \right]$$

$$M = \frac{S_i(1 - S_i)}{\text{var } S_i}$$



Summary of participation period

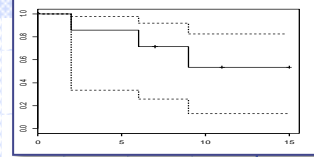
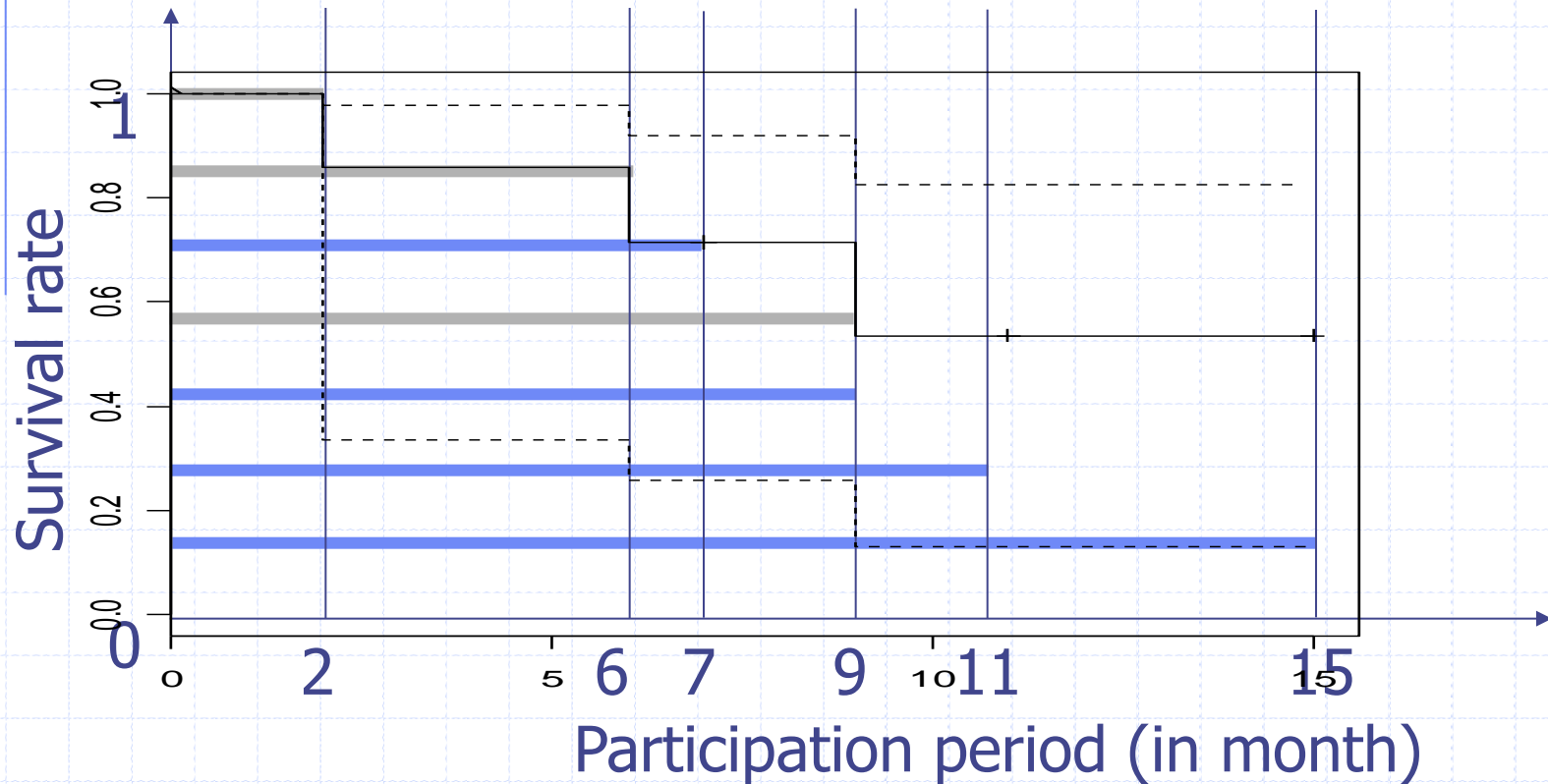




Kaplan-Meier method

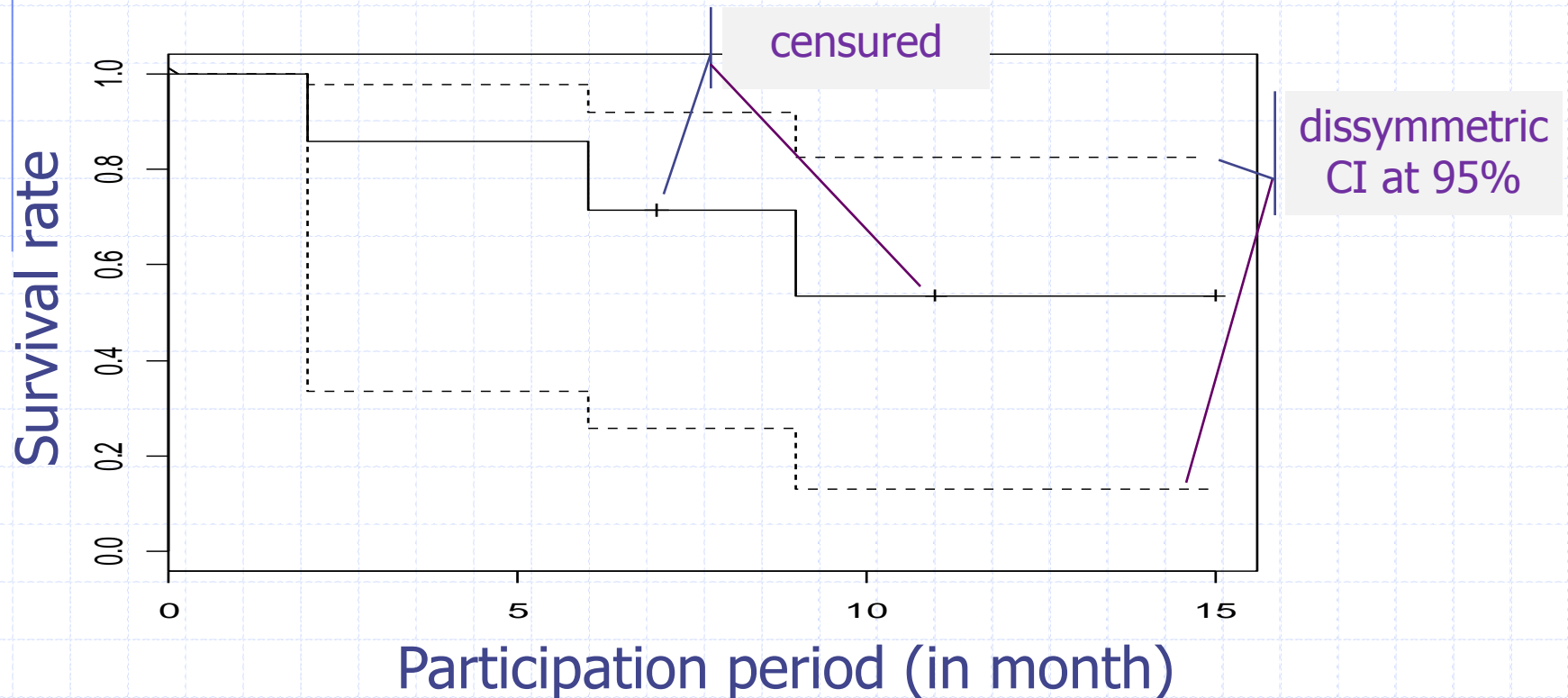
t_i	$[t_i, t_{i+1}[$	N_i	D_i	C_i	$S_{i i-1}$	S_i	$\text{var}S_i$
0	$[0, 2[$	7	0	0	1	1	-
2	$[2, 6[$	7	1	0	$6/7=0.857$	0.857	0.017
6	$[6, 7[$	6	1	0	$5/6=0.833$	$0.833 \times 0.857 = 0.714$	0.029
7	$[7, 9[$	5	0	1	1	0.714	0.029
9	$[9, 11[$	4	1	1	$3/4=0.75$	$0.714 \times 0.75 = 0.536$	0.040
11	$[11, 15]$	2	0	1	1	0.536	0.040

Graphic representation of the survival curve (Kaplan-Meier)

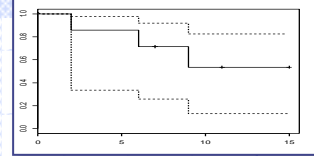


Graphic representation of the survival curve (Kaplan-Meier)

```
library(survival)
plot(survfit(Surv(PP, dead) ~ 1, conf.type="log-log"))
```



Graphic representation of the survival curve (Kaplan-Meier)



```
library(survival)
plot(survfit(Surv(PPn,d11$dead)~1,conf.type="log-log"))
```

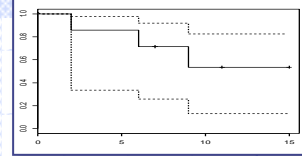
Participation period

Dichotomic variable:
dead/alive status
at the end of the par

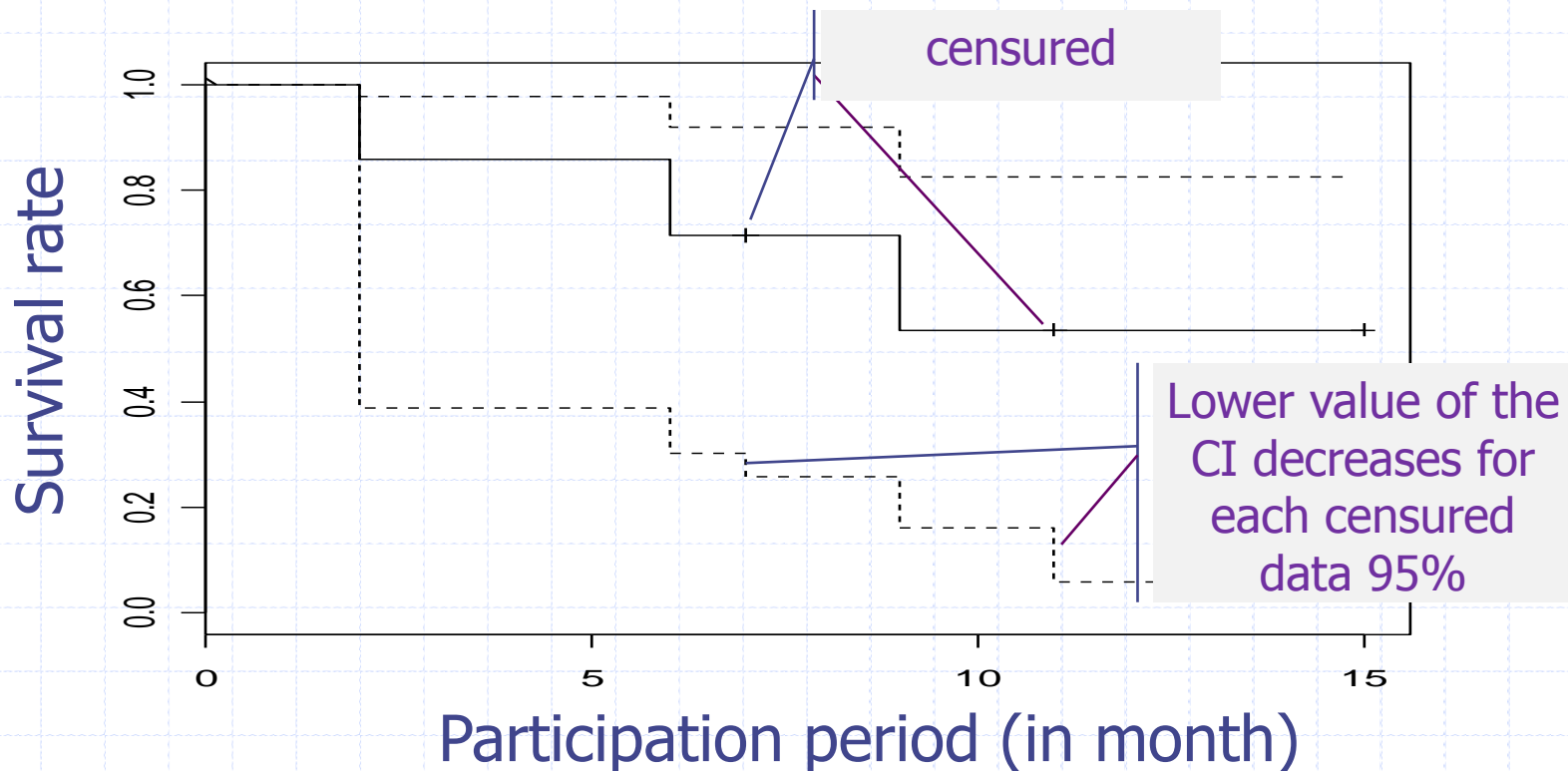
`survfit()` of the package `survival` computes all the value of the survival curve

The package `survival` need to be previously download on the computer and declare it.

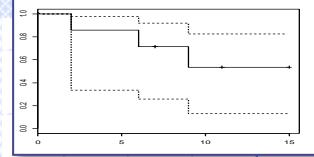
Graphic representation of the survival curve (Kaplan-Meier)



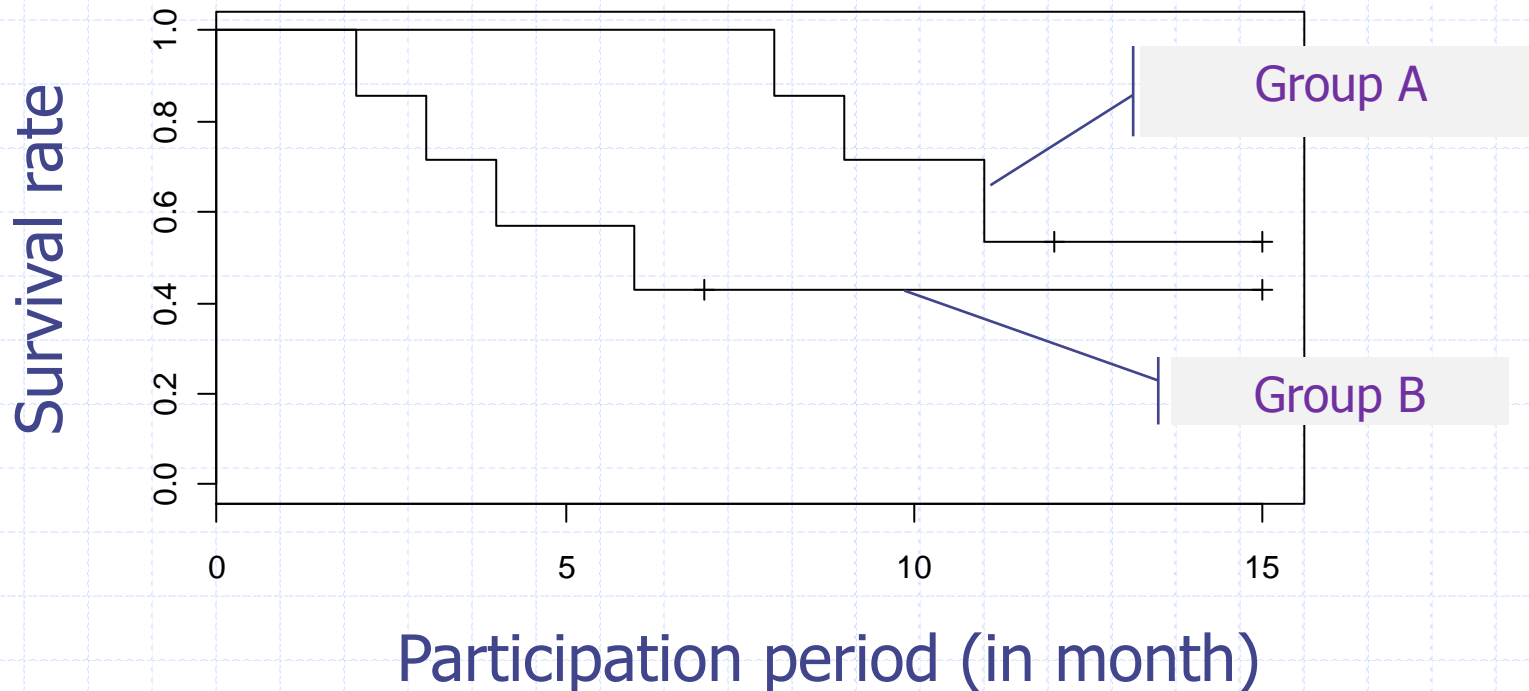
```
plot(survfit(Surv(PP, dead)~1, conf.type="log-log", conf.lower="modified"))
```

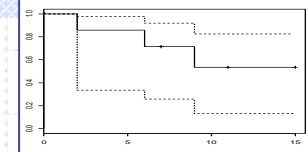


Graphic representation of survival curves for different groups



```
plot(survfit(Surv(PP, dead) ~ group, conf.type="log-log"))
```





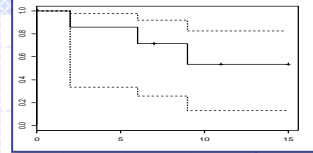
Practical on survival curve

A cohort clinical study follows more than 600 HIV sero-positive people, it is noted all the information about the date of origin (DO), the date of death (DCD).

a) Calculate the participation period for every people.

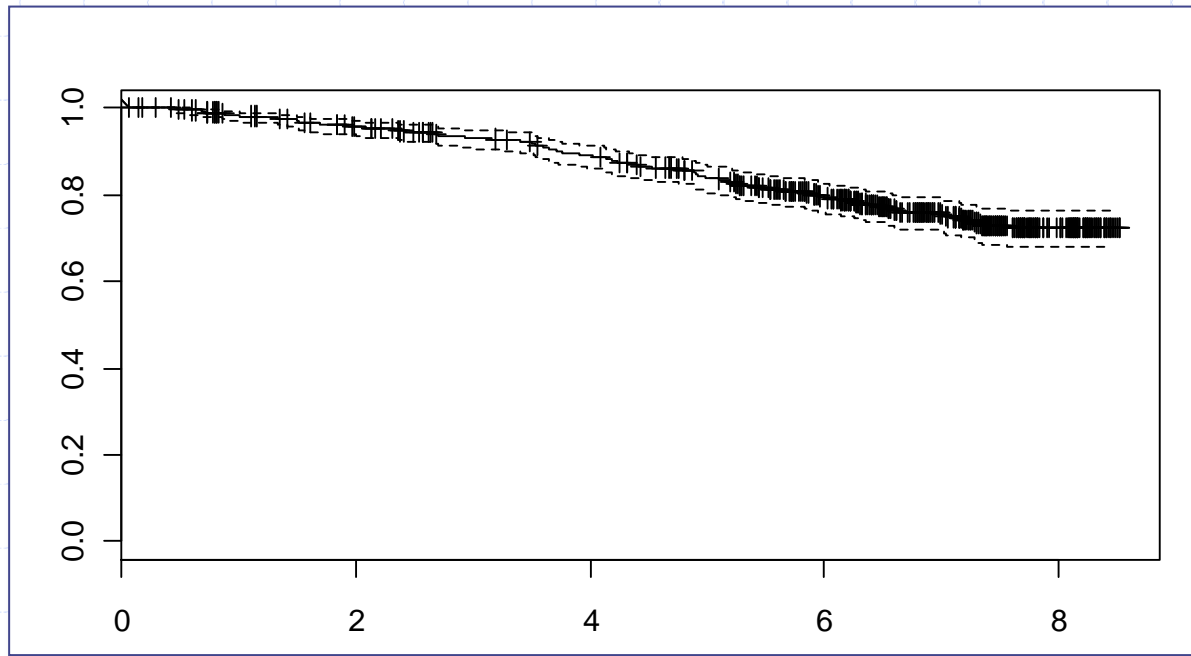
```
d10$DLNn<-as.Date(d10$DLN,"%d/%m/%Y")
d10$DIn<-as.Date(d10$DCD,"%d/%m/%Y")
d10$DOn<-as.Date(d10$DO,"%d/%m/%Y")
d10$DEn<-pmin(d10$DLNn,d10$DIn,na.rm=T)
PP<-d10$DEn-d10$DOn
PPn<-as.numeric(PP)
```

Practical on survival curve



b) Draw the survival curve.

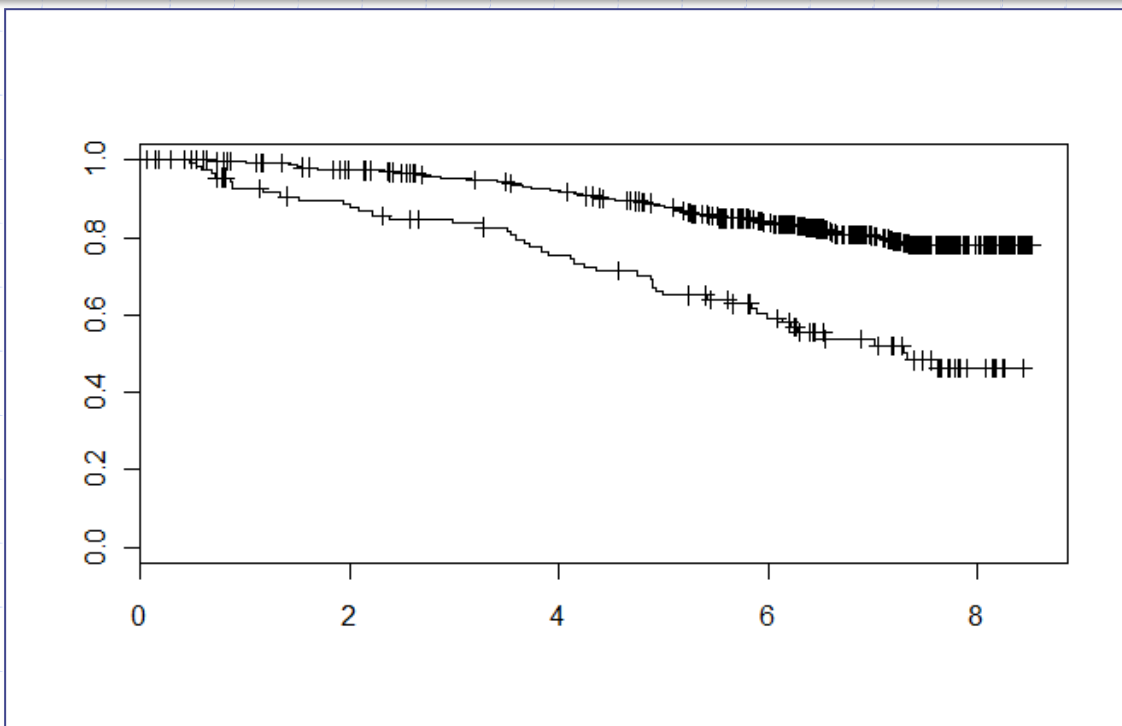
```
plot(survfit(Surv(PPn,d10$dead)~1,conf.type="log-log"))
```

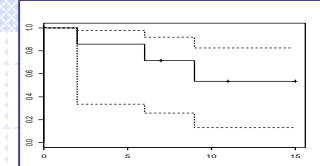


Practical on survival curve

c) Draw the survival curve according the categorie "AGE_CLASS".
AGE_CLASS=0 means younger than 40; AGE_CLASS=1 means equal or older than 40.

```
plot(survfit(Surv(PPn,d10$dead)~d10$AGE_CLASS,conf.type="log-log"))
```





Practical on survival curve

d) Read all the information about the survival curves.

```
summary(survfit(Surv(PPn,d10$dead)~d10$AGE_CLASS,conf.type="log-log"))
```

```
Call: survfit(formula = Surv(PP, d10$dead) ~ d10$AGE_CLASS, conf.type = "log-log")
```

d10\$AGE_CLASS=0

time	n.risk	n.event	survival	std.err	lower 95% CI	upper 95% CI
0.698	486	1	0.998	0.00206	0.985	1.000
0.997	479	1	0.996	0.00292	0.984	0.999
1.005	478	1	0.994	0.00358	0.981	0.998
1.333	474	1	0.992	0.00414	0.978	0.997
1.418	472	1	0.990	0.00464	0.975	0.996
1.473	471	1	0.987	0.00508	0.972	0.994
...						



To Finish....

.... Learn more

Learn more

- ◆ In R software:

?name of the function

Ex.

? cor.test

- ◆ In web:

- R cran: <http://www.r-project.org/>
- Statistics with R of Vincent Zoonekynd: google + statistics with R

- ◆ With pdf documents:

- R for Beginners of E. Paradis: google + R for beginners

- ◆ With Books:

- There is now a lot of books on R and ...
- ... there is the “Bible”: the R book of Michael Crawley

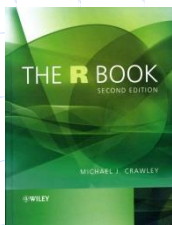
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Learn more and get help

Look at the help on chi-squared test on R program.

chisq.test {stats}

R Documentation

Pearson's Chi-squared Test for Count Data

Description

`chisq.test` performs chi-squared contingency table tests and goodness-of-fit tests.

Usage

```
chisq.test(x, y = NULL, correct = TRUE,  
           p = rep(1/length(x), length(x)), rescale.p = FALSE,  
           simulate.p.value = FALSE, B = 2000)
```

Arguments

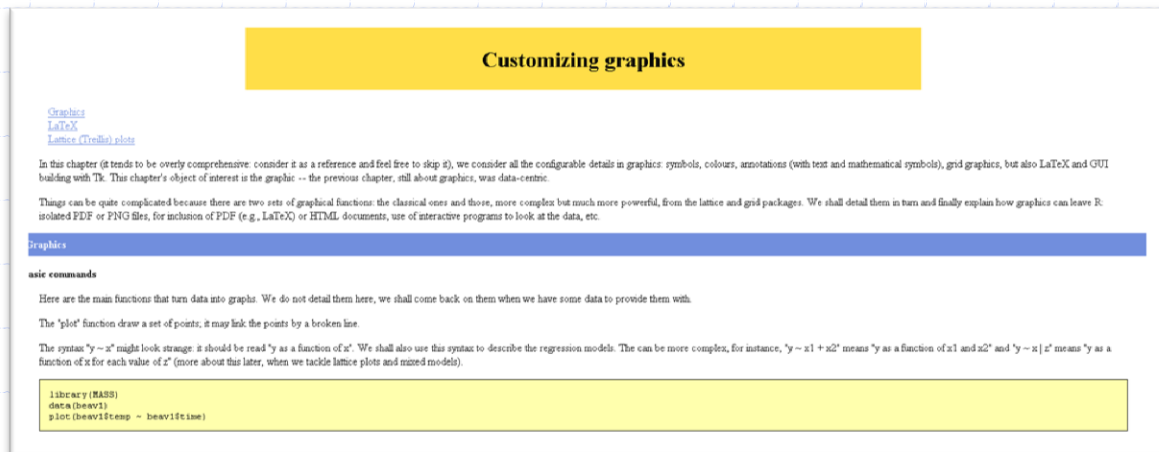
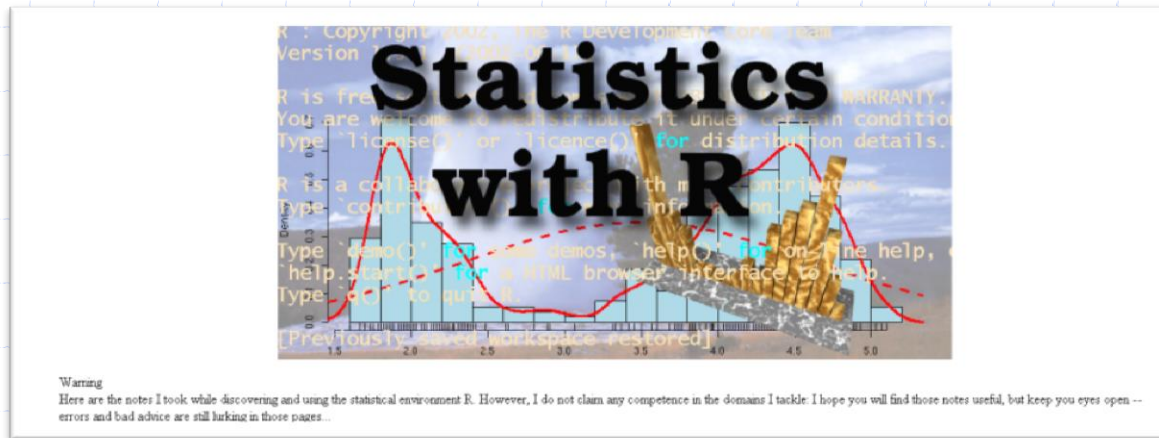
x	a numeric vector or matrix. x and y can also both be factors.
y	a numeric vector; ignored if x is a matrix. If x is a factor, y should be a factor of the same length.
correct	a logical indicating whether to apply continuity correction when computing the test statistic for 2 by 2 tables: one half is subtracted from all $ O - E $ differences; however, the correction will not be bigger than the differences themselves. No correction is done if <code>simulate.p.value = TRUE</code> .
p	a vector of probabilities of the same length of x . An error is given if any entry of p is negative.
rescale.p	a logical scalar; if TRUE then p is rescaled (if necessary) to sum to 1. If <code>rescale.p</code> is FALSE, and p does not sum to 1, an error is given.
simulate.p.value	a logical indicating whether to compute p-values by Monte Carlo simulation.
B	an integer specifying the number of replicates used in the Monte Carlo test.

Details

If **x** is a matrix with one row or column, or if **x** is a vector and **y** is not given, then a *goodness-of-fit test* is performed (**x** is treated as a one-dimensional contingency table). The entries of **x** must be non-negative integers. In this case, the hypothesis tested is whether the population

Learn more and get help

Look for the customizing graphic section in the website statistics with R



Learn more and get help

Make a research on risk ratio in the website of R cran project.

Google search results for "rate ratio site:r-project.org".

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Environ 3 400 résultats (0,25 secondes)

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cran.r-project.org/web/packages/rateratio.test/rateratio.test.pdf ▾ Traduire cette page
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[PDF] **Test Ratio of 2 Poisson Rates**
cran.r-project.org/web/packages/rateratio.test/rateratio.test.pdf ▾ Traduire cette page
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Testing the **Ratio** of Two Poisson **Rates**. Michael P. Fay. January 22, 2014. 1 Example. Here is a quick example of the function rateratio.test. Suppose you have ...

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23 déc. 2004 - observed, expected, standardized incidence **ratio**, and confidence interval. \$**rate** crude.**rate**, adjusted **rate**, and confidence interval. Note.

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